

Numerical Experiments with AMLET, a New Monte-Carlo Algorithm for Estimating Mixed Logit Models

Monte-Carlo methods with variable sampling size for mixed logit

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PART I: Theoretical investigations

- Introduction: the mixed logit problem
- Properties
 - Convergence of solutions
 - Asymptotic properties: bias and error estimation
- Estimation algorithm
 - Trust region methods
 - Variable sampling size technique

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Outline (2)

PART II: Numerical investigations

- The software AMLET
- Optimization framework: linesearch vs trust-region
- Tests on simulated data
- Comparison with Halton sequences
- Tests on real data (MobiDrive)
- Conclusions and research perspectives







Discrete choice models

Individual I

alternatives

. . .

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- Properties
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- Variable size
- AMLET
- Simulated data
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- Real data
- Conclusions

Set of alternatives available for individual *i*: $\mathcal{A}(i)$. Utility U_{ij} of $A_j \in \mathcal{A}(i)$: $U_{ij} = V(\beta) + \epsilon_{ij}$.

 β : parameters to be estimated.







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 if $U_{ij} \ge U_{in}, \forall A_n \in \mathcal{A}(i)$.





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Gumbel distributed residuals ϵ_{ij} (mean 0, scale factor μ): multinomial logit (MNL).

Probability that individual *i* choose A_j :

$$P_{ij} = \frac{e^{\mu V_{ij}(\beta)}}{\sum_{n=1}^{N} e^{\mu V_{in}(\beta)}}$$

UII (Q)



Allow heterogeneity in parameters inside the population.

$$oldsymbol{eta}=eta(oldsymbol{\gamma}, heta)$$
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 θ : vector of parameters, e.g. vector of means and std dev.



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 γ : random vector, e.g. vector of independent N(0,1);

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$$P_{ij}(\theta) = E_P \left[L_{ij}(\gamma, \theta) \right] = \int L_{ij}(\gamma, \theta) f(\gamma) d\gamma$$





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Sample average approximation (SAA) problem:

$$\max_{\theta} \hat{g}_R(\theta) = \max_{\theta} SLL(\theta) = \max_{\theta} \frac{1}{I} \sum_{i=1}^{I} \ln SP_{ij_i}^R$$



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Numerical issues

Reduction of numerical cost:

- 1. work on the objective form: quasi Monte-Carlo;
 - 2. adapt the optimization method: little exploration.







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Requirement: ability to estimate the approximation's accuracy.



Analogy to stochastic programming:

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Analogy to stochastic programming:

$$\min_x g(x) = \min_x E[f(x, \boldsymbol{\omega})]$$

 $\max_{\theta} g\left(\theta\right) = \max_{\theta} LL(\theta) = \frac{1}{I} \sum_{i} \ln E_P \left[L_{ij_i}\left(\boldsymbol{\gamma}, \theta\right) \right]$

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Mixed logit:

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Known properties can be adapted. Assume I fixed and R grows toward ∞ .





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Assume I fixed and R grows toward ∞ .

If θ_R^* , $R = 1, \ldots$, is first-order critical for the corresponding SAA problem, any limit point θ^* of $(\theta_R^*)_{R=1}^{\infty}$ is first-order critical, almost surely.



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Error and bias of simulation

With an i.i.d. sample for each individual, we have, from the delta method (see Shapiro and Rubinstein):

$$LL(\theta) - SLL^{R}(\theta) \Rightarrow N\left(0, \frac{1}{I}\sqrt{\sum_{i=1}^{I} \frac{\sigma_{ij_{i}}^{2}(\theta)}{R(P_{ij_{i}}(\theta))^{2}}}\right)$$







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Asymptotic value of the confidence interval radius:

$$\epsilon_{\delta} = \alpha_{\delta} \frac{1}{I} \sqrt{\sum_{i=1}^{I} \frac{\sigma_{ij_i}^2(\theta)}{R(P_{ij_i}(\theta))^2}}$$

 δ : signification level; $\alpha_{0.9} \approx 1.65$.





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$$B := E[SLL^{R}(\theta)] - LL(\theta) = -\frac{I\epsilon_{\delta}^{2}}{2\alpha_{\delta}^{2}}$$

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In practice, use of SAA estimators $\sigma_{ij_i}^R(\theta)$ and $P_{ij_i}^R(\theta)$.

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Trust-region methods

Basic principle: at iteration k, k = 1, ..., approximately minimize a model of the objective over a trust region (\mathcal{B}_k). It gives a (candidate iterate).





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Compute the following ratio:

$$\rho = \frac{\rm real \ reduction}{\rm predicted \ reduction}$$

• If $\rho \ge \eta_1$, accept the candidate.

If $\rho \geq \eta_2$, enlarge \mathcal{B}_k , otherwise keep it the same, or reduce it.

• If $\rho < \eta_1$, reject the candidate and reduce \mathcal{B}_k .

For instance,
$$\eta_1 = 0.01$$
 and $\eta_2 = 0.75$.



Example (Conn, Gould, Toint 2000):

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$$\min_{x,y} -10x^2 + 10y^2 + 4\sin(xy) - 2x + x^4$$



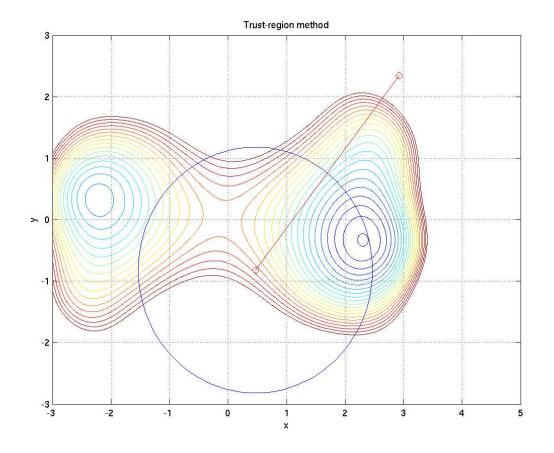


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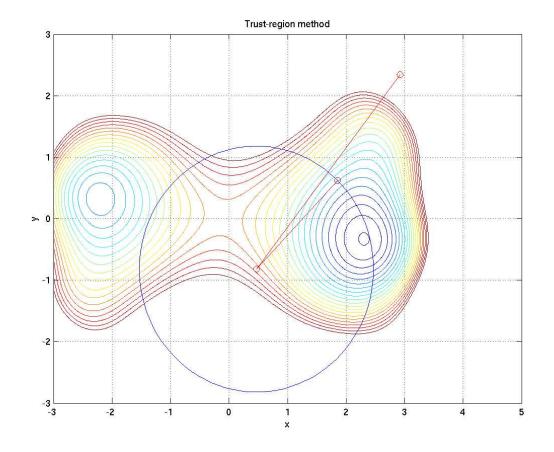


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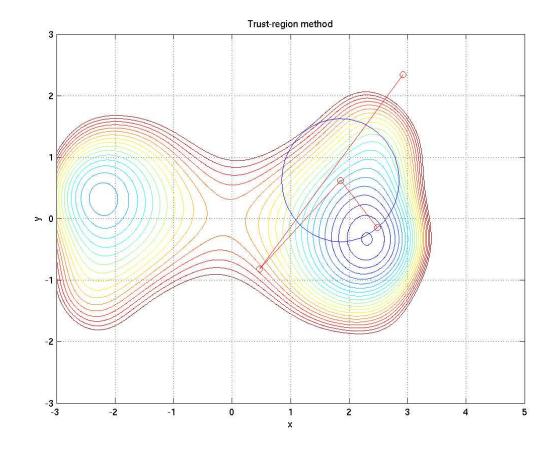


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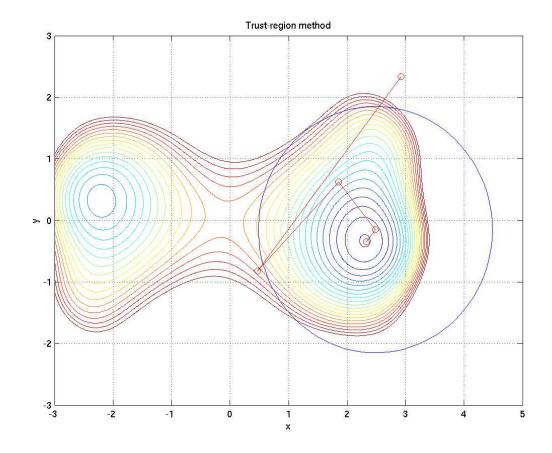


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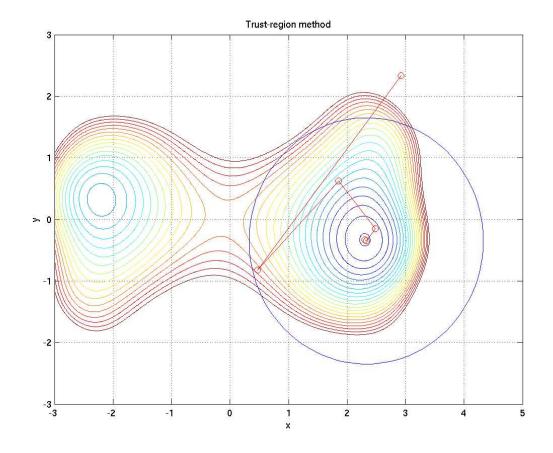


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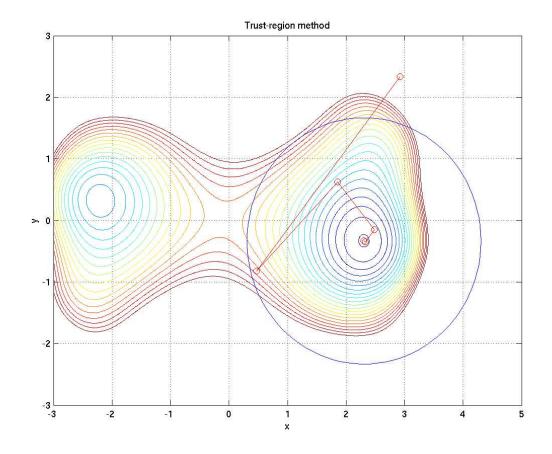


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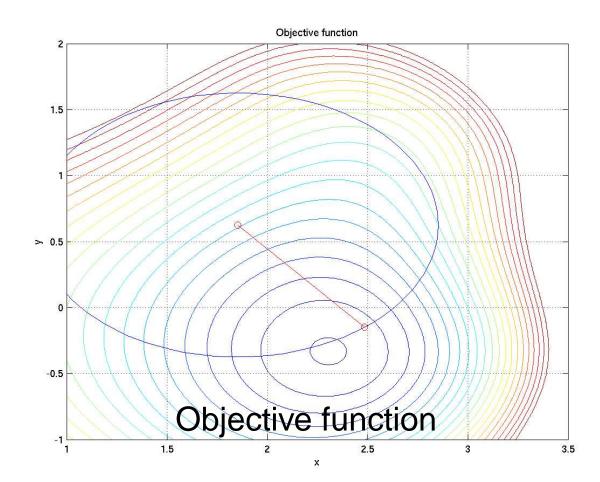




Model and objective comparison

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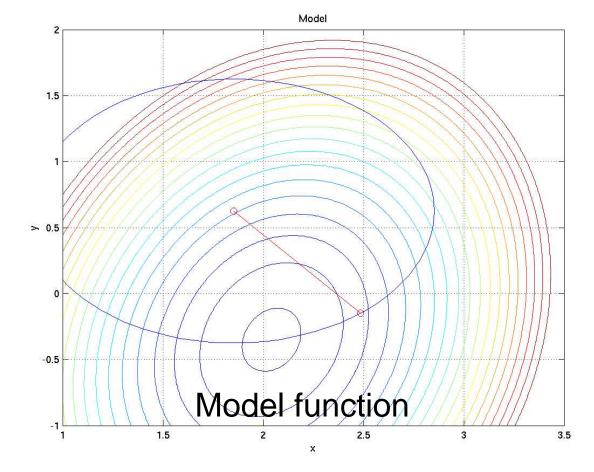




Model and objective comparison

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Variable sampling size technique

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New software: AMLET (Another Mixed Logit Estimation Tool)









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New software: AMLET (Another Mixed Logit Estimation Tool)

 multinomial logit and mixed logit models support;









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- written in C, tested on a Linux system.

Working environment: Pentium IV, 2Ghz.





Tests on simulated data

Experimental design:

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Tests on simulated data

Experimental design:

- attribute values drawn drom a N(0,1);
- random parameters $\sim N(0.5, 1)$;
- linear utilities;
- number of alternatives: 2, 3, 5, 10;
- observations: 2000, 5000, 7500, 1000.

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Comparisons with Gauss 5.0 with MaxLik modules, and code of Train (Halton, tol=10⁻⁶). Note: Gauss tested on Pentium III (1 licence); reported times are accordingly corrected.



Linesearch versus Trust-Region

Is the trust-region choice appropriate?

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Linesearch versus Trust-Region

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Trust-regionVariable size		abel linesearch (STEPBT)		
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HaltonReal data		mendation of Nocedal and Wright)		
 Conclusions 		Basic Trust-Region (BTR)		
		BTR with dynamic accuracy (BTRDA)		







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Simulation experiment:

> 5000 individuals, 5 alternatives, 5 independent random parameters (N(0.5, 1.0));

 \triangleright minimization over 10 random samples of size 2000.

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Linesearch versus Trust-Region (2)

	Set	BFGS		BTR		BTRDA	
ine		Likelihood	Time (s)	Likelihood	Time (s)	Likelihood	Time (s)
oduction perties	1	-1.40529	1421	-1.40529	1103	-1.40529	722
st-region iable size	2	-1.40532	1455	-1.40532	997	-1.40532	618
LET	3	-1.40519	1691	-1.40519	1090	-1.40519	730
<i>imulated data</i> ton	4	-1.40409	1736	-1.40409	997	-1.40409	564
al data nclusions	5	-1.40525	1702	-1.40525	1017	-1.40525	531
	6	-1.40499	1707	-1.40499	1089	-1.40499	587
	7	-1.40437	1475	-1.40437	1045	-1.40437	560
	8	-1.40530	1706	-1.40530	998	-1.40530	781
	9	-1.40532	1701	-1.40532	1028	-1.40532	724
	10	-1.40441	1729	-1.40441	1044	-1.40441	546
	Mean	-1.40495	1632	-1.40495	1040	-1.40495	636



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▷ Clear advantage to the trust-region approach.





Tests on simulated data

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Comparison with Halton sequences

Halton sequences are popular: allow smaller sizes.

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Halton sequences are popular: allow smaller sizes. Comparisons with Gauss for low dimensions (≤ 5 random parameters): similar results, but AMLET remains faster (2 to 5 times for 1000 Monte-Carlo, compared to 125 Halton).







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Halton sequences are popular: allow smaller sizes. Comparisons with Gauss for low dimensions (≤ 5 random parameters): similar results, but AMLET remains faster (2 to 5 times for 1000 Monte-Carlo, compared to 125 Halton). High dimensions with Halton sequences: loss of uniform coverage. Graphs with 250 Halton numbers:

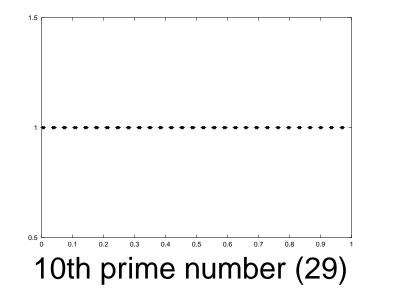


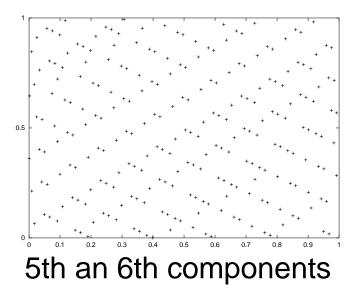


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Halton vs Monte-Carlo

· MAMUR. T	Variable	125 Halt.	250 Halt.	1000 MC	2000 MC	3000 MC
	P1 mean	0.4022	0.4371	0.437926	0.451574	0.452668
	P1 std. dev.	0.9575	1.0546	1.06785	1.10345	1.10275
	P2 mean	0.4237	0.4586	0.460888	0.47644	0.476496
Outline	P2 std. dev.	0.8423	0.8646	0.889336	0.924625	0.929872
 Introduction 	P3 mean	0.3903	0.4305	0.428926	0.442901	0.444273
 Properties 	P3 std. dev.	0.7959	0.8812	0.881386	0.934353	0.934465
 Trust-region 	P4 mean	0.4700	0.5129	0.51648	0.533209	0.534616
 Variable size 	P4 std. dev.	0.6744	0.7205	0.744546	0.77631	0.780702
• AMLET	P5 mean	0.4808	0.5308	0.531721	0.549948	0.5521
 Simulated data 	P5 std. dev.	0.7027	0.8312	0.847629	0.883452	0.886892
ightarrow Halton	P6 mean	0.3782	0.4100	0.412724	0.425807	0.426386
 Real data 	P6 std. dev.	0.8297	0.9163	0.947523	0.987107	0.986711
 Conclusions 	P7 mean	0.3920	0.4337	0.435504	0.450424	0.450725
	P7 std. dev.	0.9116	1.0398	1.03509	1.07814	1.0786
	P8 mean	0.4895	0.5310	0.534592	0.551502	0.552986
	P8 std. dev.	0.9198	1.0022	1.02172	1.054758	1.06168
	P9 mean	0.4441	0.4768	0.481642	0.498321	0.49854
	P9 std. dev.	0.9048	0.9694	0.977834	1.02296	1.02425
	P10 mean	0.3702	0.4149	0.412852	0.426316	0.426021
	P10 std. dev.	0.6250	0.8482	0.808399	0.835485	0.840841
	Log-likelihood	-1.44244	-1.43990	-1.44132	-1.44086	-1.44077
	Accuracy	NA	NA	0.001032	0.0007402	0.0006049
The	Bias	NA	NA	-0.0009844	-0.0005062	-0.0003381
University of Namur	Time (s)	2077	4252	435.8	890	1349.1

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Halton vs Monte-Carlo: remarks

• AMLET very fast compared to Gauss;

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- 250 Halton give similar solution to 1000 Monte-Carlo, but results change with 2000 Monte-Carlo, and Gauss underevaluates the objective;







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Quid of Halton sequences? Our results incite us to be careful concerning their usage, but we still guess that good results can be obtained with scrambled sequences (see Bhat).

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Tests on real data

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Mode choice model: Mobidrive data (Axhausen and al.)

- \rightarrow 5799 observations;
- \rightarrow 5 alternatives;
- \rightarrow 3 random parameters (normally distributed).







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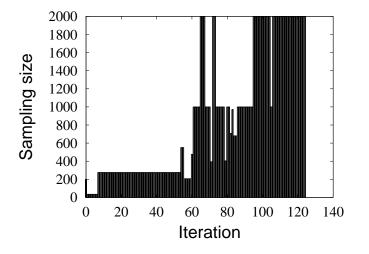
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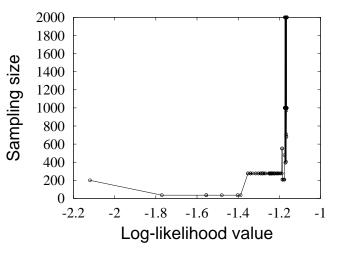
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Evolution of sample sizes (2000 Monte-Carlo)





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Tests on real data: results

	Variable	Gauss (125 Halt.)	<u>АМLЕТ</u> (1000 MC)	AMLET (2000 MC)		
	Car Passenger (CD)	-1.4511	-1.45104	-1.4527		
Outline	Public Transport (PT)	-0.9355	-0.932594	-0.932458		
 Introduction 	Walk (W)	0.1081	0.109186	0.108793		
 Properties 	Bike (B)	-0.6355	-0.634269	-0.635217		
 Trust-region 	Urban household location (PT)	0.560609	0.557866	0.561515		
• Variable size	Suburban household location (W, B)	-0.3451	-0.345403	-0.345113		
• AMLET	Full-time worker (PT)	0.2690	0.269265	0.268996		
 Simulated data 	Female and part-time (CP)	0.9133	0.912925	0.913835		
● Halton → Real data	Married with children (CD)	0.9716	0.970755	0.971656		
 Conclusions 	Annual mileage (CD)	0.0518	0.0518679	0.0519161		
	Number of stop (CD)	0.1349	0.135187	0.135817		
	Time mean	-0.0268	-0.0268882	-0.0269985		
	Time std. dev.	0.0205	0.0206265	0.0208197		
	Cost mean	-0.1683	-0.168923	-0.169365		
	Cost std. dev.	-0.0452	0.0465829	0.0465628		
	Time budget/100 mean (CD, CP)	-0.1249	-0.124816	-0.125128		
	Time budget/100 std. dev. (CD, CP)	-0.1136	0.112801	0.113803		
	Log-likelihood	-1.16489	-1.16479	-1.16470		
The	Bias	Not available	-0.00009117	-0.0000463		
University ^{of} Namur	Accuracy	Not available	0.0002916	0.0002086		
	Time (s)	2439	936	1549		
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 New Monte-Carlo algorithm for nonlinear stochastic programming;







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Conclusions

- New Monte-Carlo algorithm for nonlinear stochastic programming;
- Estimation of error and bias in the objective
 - allow variable sampling size techniques;
 - not sufficient to know stability of solutions;





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Conclusions

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- Estimation of error and bias in the objective
 - allow variable sampling size techniques;
 - not sufficient to know stability of solutions;
- Behaviour seems to be better when the number of alternatives rises;
- Implementation issues are important:
 - choice of optimization method;
 - implementation tricks (e.g.: possible to compute objective and gradient simultaneously).



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Perspectives

- AMLET still on a prototype level:
 - more tests on real data would be useful;
 - planned features:
 - other distributions than normals;
 - correlation between observations;





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Perspectives

- AMLET still on a prototype level:
 - more tests on real data would be useful;
 - planned features:
 - other distributions than normals;
 - correlation between observations;
- variable sampling size algorithm could be certainly refined;
- investigation of quasi Monte-Carlo methods:
 - exploitable error bounds?
 - exploration of scrambled sequences;
 - could reduce computation time and memory needs.