

# Buy As You Need: Nutrition and Food Storage Imperfections

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February 2017

## Abstract

In this paper, we investigate how the activation of local food markets impacts the nutritional status of both children and adults, in a context characterized by large seasonal fluctuations in the price and availability of foodgrain. Taking advantage of the random scaling-up of a program of Food Security Granaries (FSGs) in Burkina Faso, we reach three conclusions. First, especially in remote areas where local markets are thin, food market activation considerably dampens nutritional stress. The effect is strongest among children, and young children in particular, for whom deficient nutrition has devastating long-term consequences. Second, and surprisingly, this beneficial effect is obtained despite the fact that total food consumption does not increase as a result of the external intervention. Third, it is a change in the timing of food purchase, translated into a change in the timing of consumption, that drives the nutritional improvement. A simple two-period model shows that an increase in consumption needs not take place when the price of foodgrain declines during the lean season if storage losses are taken into account. More than the waste of the foodgrain stored, it is the urge to consume purchased foodgrain which gives rise to storage imperfections: foodgrain purchased in anticipation of uncertain future supply results in immediate consumption and body mass accumulation, which is less efficient than nutrition-smoothing consumption flows.

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<sup>†</sup>This study would not have been possible without the logistical and financial support of SOS Faim. We acknowledge funding by the European Union under the 7th Research Framework Programme (theme SSH) - Grant agreement nr 290752 (Nopoor project). Our gratitude also extends to participants in various seminars or conferences organized by NEUDC (2016), CSAE (2016), UCDavis, Cornell University and the Catholic University of Louvain. Special thanks are due to Jean-Marie Baland, Angus Deaton, Dilip Mookherjee, Kalle Hirvonen and William Pariente who has commented upon several successive versions of this paper.

# 1 Introduction

Food insecurity remains a serious problem today. In sub-Saharan Africa, in particular, 40 percent of all children are stunted and more than 20 percent are underweight (Black et al., 2013), with disastrous consequences for their long term development (Glewwe et. al., 2001, Alderman et. al., 2004). If we leave fragile states aside, food insecurity tends to be concentrated in remote areas of the developing world. By definition, these areas have scattered populations that are physically isolated and are characterized by pervasive market imperfections associated with high transaction costs (see for example, De Janvry et al., 1991, Fafchamps, 1992, Renkow et al., 2004). In this paper, we investigate how the activation of local food markets impacts the nutritional status of both children and adults, in a context characterized by large seasonal fluctuations in the price and availability of foodgrain. By market activation, we mean the creation of a more steady local supply of foodgrain throughout the agricultural cycle. Because we carried out the impact evaluation in a bad agricultural year, we are able to detect clear effects and also to highlight the nature of the most binding constraints confronting rural households in search for better nutrition.

According to *prima facie* economic intuition, the intended effects of the program are : (1) an increased consumption of foodgrain in response to lower average prices (the income effect), and (2) a smoother distribution of food purchases over the year in response to a more regular supply (the availability effect). Our study shows that while the second effect is confirmed, the first one is not. To elucidate these two results together, we need to distinguish between food consumption and nutrition. This is done in the framework of a simple model inspired by Deaton and Muellbauer (1980) and Dercon and Krishnan (2000a) in which household utility depends on nutrition measured as a stock. We show that a greater availability of foodgrain on the local market may enhance nutrition without necessitating an increase in consumption. This is the result of the fact that households are now able to avoid costly storage, whether in the form of foodgrain or body mass.

One of the most original aspects of our work is that the household stock management strategies can be apprehended in relation to market purchases, exploiting the fact that the degree of market activity varies randomly across villages. Toward that end, we have collected exceptionally rich data including the nutritional status of all household members, detailed information about foodgrain transactions and stock availability at the household level. Moreover, we rely on qualitative investigation tools that we have especially designed for the purpose of testing our quantitative evidence and our interpretation of the main findings. These tools go beyond simple group discussions to include interactive activities in

which a restricted number of participants were invited to make critical intertemporal allocation choices and explain them in detail.

We show that the large improvement in nutrition resulting from market activation is not caused by an increase in total food purchases and consumption but by delayed purchases that allow a more efficient allocation of consumption across the year. The implication is that the timing of purchases and the timing of consumption are not independent of each other as would be the case if there were no savings and storage imperfections. Storage imperfections under the form of pressure to consume purchased foodgrain quickly incite households to minimize storage and postpone purchases to times of need, provided that future supply is better assured.

We contribute to several strands of the literature. First there is a literature dealing with issues of stock management and savings at the family level yet it largely focuses on the self-insurance property of stocks to confront unpredictable shocks (see the pioneering works of Newbery, 1991: 284-8 and Platteau, 1991). Moreover, this literature tends to illustrate stock management strategies by reference to livestock (Fafchamps et al., 1998). In our study, we are concerned with foodgrain stocks and with their management in a context of anticipated seasonal shortages rather than unpredictable shocks. A theoretical literature also exists that analyzes optimal household behavior in the presence of seasonality and risk aversion or credit market imperfections (Park, 2006, Stephens and Barrett, 2011). In the narrow empirical literature on this topic, a salient issue is the impact of seasonality on current consumption. There is no clear consensus emerging from it: while some studies conclude that consumption is largely smoothed over the agricultural cycle, others point to the opposite conclusion (see, for example, Paxson, 1993, for the former conclusion, and Dercon and Krishnan, 2000b, for the latter).

Still another strand of literature uses anthropometric measures instead of consumption with a view to assessing fluctuations in the nutritional level of individual household members. Attention is usually given to the determinants and consequences of these fluctuations (Dercon and Krishnan, 2000a, Alderman et al., 2006, Vaitla et al., 2009) and to the short-term and long-term effects of various policies designed to combat malnutrition (Hoddinott et al., 2008, Yamano et al., 2005). Closely related to our endeavour, Abay and Hirvonen (2016) examine the seasonal weight fluctuations of young children in Ethiopia and provide evidence that market integration dampens children's exposure to seasonal shortages. They suggest that diet diversity may explain the observed patterns.

Finally, a relevant but restricted set of studies queries about the effectiveness and sustainability of cereal banks, which are community-based interventions aimed at stabilizing food supply across the

year. While there is limited evidence of their impact (Basu and Wong, 2015, Barrett, 1996), cereal banks and their derivatives have benefited from a resurgence of interest over the last decade. The World Food Program, the European Union, Non-Governmental Organizations and local authorities, have started again to fund thousands of initiatives designed to promote food security through the building of local food reserves in Sahelian countries (Oxfam International, 2013, World Bank, 2012).

The outline of the paper is as follows. In Section 2, we present a simple two-period model adapted from Dercon and Krishnan (2000a), to investigate how a decrease in the real price of foodgrain in the second (lean) period affects consumption behavior and nutritional outcomes in both periods. In Section 3, we describe the context of our experiment and the data. In Section 4, we discuss our empirical strategy before estimating the impacts of market activation on food access, purchase, consumption and nutrition. Section 5 is devoted to the interpretation of the results and draws on a variety of qualitative and quantitative evidence gathered *ex post*. Section 6 concludes.

## 2 Model

We propose a simple framework to analyze the household's problem of allocating food consumption across two periods that follow a single harvest.

We consider a household whose utility depends on its nutritional status in period  $t$ ,  $N_t$ , and the consumption of a numeraire,  $O_t$ . It is written as:  $U(N_t, O_t)$ . There are two periods: a dry, post-harvest season ( $t = 1$ ) succeeded by a rainy, lean season ( $t = 2$ ). The household enters period 1 with a nutritional status  $N_0$ . Its problem is to intertemporally allocate food consumption to maximize:

$$U(N_1, O_1) + \gamma U(N_2, O_2)$$

where  $\gamma$  is a discount factor.

We follow Dercon and Krishnan (2000a) by modelling the nutritional status as a stock or durable, and specifying the nutritional status in each period as:

$$N_1 = f(N_0) + n(C_1) - a(L_1) \tag{1}$$

$$N_2 = f(N_1) + n(C_2) - a(L_2) \tag{2}$$

These equations indicate that the nutritional status in each period is a function of the nutritional status in the previous period, where the function  $f$  captures the depreciation of the nutrition stock between periods, with  $0 < f' < 1$ . The nutritional status also increases with current period consumption ( $C_1$  or  $C_2$ ), according to the transformation function  $n$ , with  $n' \geq 0$  and  $n'' < 0$ . Finally, expending energy will reduce nutritional status: this loss is taken into account via the function  $a$ , which is increasing in labour effort,  $L$ . For the sake of simplicity, we assume that  $L_1$  and  $L_2$  are given.

The household consumes both his own production fully available in period 1, denoted by  $y$ , and food bought on the market in either period 1 or 2, denoted by  $m_t$ . The food stored from period 1 to period 2 is  $s \geq 0$ . The quantity available in period 2 is then  $\beta s$ , with  $0 \leq \beta \leq 1$  to account for physical storage losses. Food availability constraints are then:

$$\begin{aligned} C_1 &= y - s + m_1 \\ C_2 &= \beta s + m_2 \end{aligned}$$

Combining the two equations and the non-negative stock constraint, we can write:

$$\beta C_1 + C_2 = \beta y + \beta m_1 + m_2 \quad (3)$$

$$C_1 \leq y + m_1 \quad (4)$$

We now need to specify the budget constraint. We assume that the household has an exogenous income  $R$ , which is obtained in period 1 only.  $R$  can be saved and will yield  $rR$  in period 2. The market price for food is  $P_1$  in period 1 and  $P_2$  in period 2, with  $P_2 > P_1$  to account for the seasonal price increase in the lean season. The budget constraints in period 1 and 2 are:

$$P_1 m_1 + b + O_1 \leq R + P_1 y$$

$$P_2 m_2 + O_2 \leq r b$$

where  $b$  is the amount of income saved in period 1.<sup>1</sup> The two constraints can be combined in a single expression:

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<sup>1</sup>The amount saved verifies:  $0 \leq b \leq R$

$$rP_1m_1 + rO_1 + P_2m_2 + O_2 \leq rR + rP_1y \quad (5)$$

We restrict attention to cases where consumption is positive in both periods and where households are net food buyers. Assuming positive consumption in both periods implies that the effectiveness of body mass storing as measured by  $f$  is never high enough to enable a household to achieve a minimum nutritional level without consuming some food during the current season. To guarantee that households are net buyers we simply set  $y$  to zero. In solving the above problem of the household, two cases arise depending on whether food purchase in period 2 is positive (Case 1) or zero (Case 2).

**Case 1:**  $P_2 < \frac{rP_1}{\beta}$

When  $P_2 < \frac{rP_1}{\beta}$ , the household buys food in period 2,  $m_2 > 0$ . For food to be purchased in period 2, the unit price in this period cannot exceed the value of one unit of food bought in period 1 and stored, taking into account the return of the money saved and the storage cost avoided when purchase is postponed (see Appendix 1).

The intertemporal allocation of nutrition across time is characterized by the following equation, where  $U_N$  is the marginal utility of nutrition (details are provided in Appendix 1):

$$\frac{U_N(N_1, O_1)}{\gamma U_N(N_2, O_2)} = \frac{\frac{1}{n'(C_1)} rP_1 - \frac{f'(N_1)}{n'(C_2)} P_2}{\frac{P_2}{n'(C_2)}} \quad (6)$$

This expression has an intuitive interpretation. It relates the intertemporal ratio of marginal utilities to the relative cost of increasing nutrition via current consumption in the first period relative to the second period. The denominator defines the marginal cost of nutrition via current consumption in period 2. The numerator is the future value of the marginal cost of nutrition via current consumption in period 1, taking into account the benefit corresponding to the carryover effect from improved nutrition in period 1 (the body mass storing effect).

If we assume that the transformation function,  $n$ , and the depreciation function,  $f$ , are linear and that the utility function is additively separable in nutrition and the numeraire (as in Dercon and Krishnan, 2001a), we can rewrite (6) as  $\frac{U_N(N_1)}{\gamma U_N(N_2)} = r\frac{P_1}{P_2} - f'$ , from which the roles of  $\frac{P_1}{P_2}$  and  $f'$  are evident. In words, when the price in the second period decreases, the ratio of marginal utilities must increase. This implies that the nutrition level in the second period will increase relative to the level in

the first period. On the other hand, if the carryover effect becomes more effective ( $f'$  increases), the opposite outcome is obtained, and nutrition in the first period increases.

**Case 2:**  $P_2 > \frac{rP_1}{\beta}$

When  $P_2 > \frac{rP_1}{\beta}$ , it follows that  $m_2 = 0$ . In this case, it is less expensive to buy a unit of food in period 1 and to store it than to buy it in period 2 for immediate consumption. The intertemporal allocation of nutrition across time is then characterized by the following equation (see Appendix 1):

$$\frac{U_N(N_1, O_1)}{\gamma U_N(N_2, O_2)} = \frac{\frac{1}{n'(C_1)} - \frac{1}{\beta n'(C_2)} f'(N_1)}{\frac{1}{\beta n'(C_2)}} \quad (7)$$

This expression has a similar interpretation as (6), except that the ratio of prices is replaced by  $\beta$ . Given that there is no purchase in period 2, the relevant cost of waiting until period 2 to consume a unit of food purchased in period 1 is measured by the inverse of the retention rate,  $\frac{1}{\beta}$ .

Assuming again linearity of the functions  $f$  and  $n$  and separability of the utility function, equation (7) can be simply rewritten as:  $\frac{U_N(N_1)}{\gamma U_N(N_2)} = \beta - f'$ . Therefore, if the storage loss increases relative to the effectiveness of body mass storing, that is, if the retention rate,  $\beta$ , decreases, the marginal utility of food in period 1 must decrease relative to the one in period 2. This implies that the nutrition level will be boosted in period 1 compared to period 2.

### The impact of the intervention

When the local market is activated, since local markets are particularly thin during the lean season, we expect  $P_2$  to decrease. The condition  $P_2 \leq \frac{rP_1}{\beta}$  is then more likely to be satisfied, implying that the household will be more likely to buy food in the second period ( $m_2 > 0$ ) and less likely to resort to body mass or household storage. Furthermore, provided  $m_2 > 0$ , a decrease in  $P_2$  will increase the nutrition level in period 2, relative to the nutrition level in period 1.<sup>2</sup> Because  $N_2$  is lower than  $N_1$  to start with, the decrease in  $P_2$  will help smoothen the nutrition level across seasons: the household will experience less fluctuation in nutrition.<sup>3</sup>

What about the effect of a decrease in  $P_2$  on food consumption? Provided that  $m_2 > 0$ , the increase in  $N_2$  relative to  $N_1$  results from an increase in  $C_2$ . Whether or not total consumption ( $C_1 + C_2$ ) increases is not clear a priori. Indeed the increase in  $C_2$  implies that the household has less

<sup>2</sup>Assuming again linearity of the functions  $f$  and  $n$  and separability of the utility function.

<sup>3</sup>We know  $N_2 < N_1$  since  $\frac{U_N(N_1)}{U_N(N_2)} = \gamma(\beta - f') < 1$

need for body mass storing and, therefore,  $N_1$  and  $C_1$  may decrease concomitantly. We illustrate the possibility that  $C_1 + C_2$  may decrease as a result of a fall in  $P_2$  with the help of a numerical simulation. Figures 6 and 7 summarize the findings obtained.

Figure 6 shows that, as expected, the relationships between  $N_1$  and  $N_2$  to  $P_2$  are both monotonous. A decrease in  $P_2$  increases  $N_2$  and decreases  $N_1$ . Figure 7 depicts the way total consumption  $C_1 + C_2$ , responds to an increase in  $P_2$ . We can see that the relationship is not monotonous and, starting from high levels of  $P_2$  (relative to  $P_1$ ), total consumption first decreases when  $P_2$  falls, and then increases.

### 3 Program and Experimental Design

#### 3.1 The Food Security Granaries program

In the late 1970s, in order to mitigate the food access problem, many aid organizations and governments have widely promoted the creation of local community organizations aimed at activating local food markets. Cereal banks are a typical example of these community-based interventions seeking to reduce market risks understood as either availability risk (food supply becomes less reliable in times of need) or price risk (food price rises in times of need). However, most of the 4000 cereal banks that were inventoried in Sahelian countries in 1991 collapsed in the late 90s owing to mismanagement, embezzlement of funds, and lack of trade opportunities (for a review of the problems, see World Bank, 2011). A new generation of initiatives inspired by the principles of cereal banks has nonetheless developed over the last decade. Foremost among them is the program of Food Security Granaries (FSG) undertaken in 2002 in Northern Burkina Faso by the NGO “SOS Faim” and financed by the Belgian Fund for Food Security (FBSA). It is aimed at revitalizing a network of about 400 former cereal banks in a setup that pays strong attention to financial viability considerations.

The northern part of Burkina Faso is of particular interest because it belongs to the Sudano-Sahelian dry zone where, given the absence of irrigation, there is only one agricultural cycle per year and production is highly sensitive to rainfall shocks. Few crops can be cultivated and the diversification of the food diet is very low: foodgrain represents more than two-thirds of daily caloric intakes (Cheyns, 1996). Subsistence agriculture dominates but while some households are systematically able to produce enough grain, most of them are not and depend crucially on food markets to satisfy their basic needs. Food access is especially critical in the rainy season when people engage in heavy agricultural work, grain stored in family granaries start depleting, food prices tend to increase, and access to villages



becomes more difficult because of rain. Hence the name lean season to characterize this period.

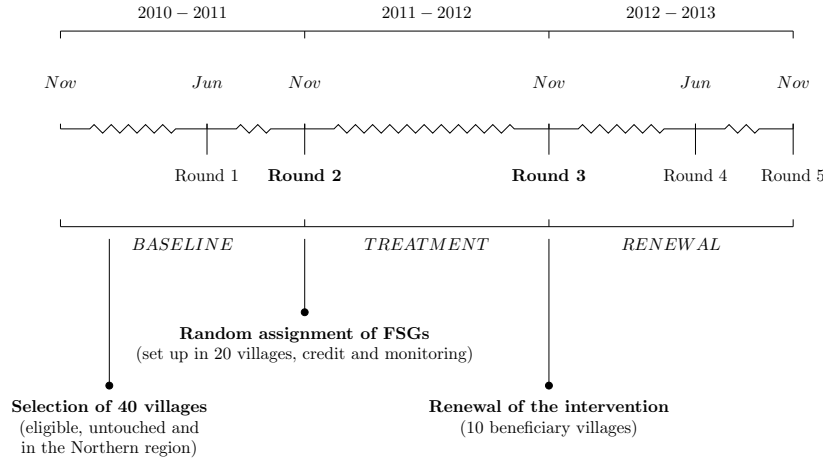
The pillars of the FSG intervention are: 1) set up a local, informal storing and marketing organization whose function is to buy foodgrain, store and sell it throughout the agricultural year, 2) provide training and capacity-building for local management teams, as well as monitoring and multi-level technical assistance on a continuous basis, 3) grant (gradually scaled up) annual credit to each village organization in the form of a revolving fund, and 4) through the network of local farmer organizations, shift grain from surplus to deficit village communities so that the latter can complement local supplies with external ones.

The last three functions are expected to provide village granaries with a comparative advantage over the private sector, thereby enabling them to operate even where and when private merchants are absent. Training and monitoring, credit-giving and foodgrain redistribution between surplus and deficit areas can be performed rather effectively thanks to the network-based operation of the scheme. As a matter of fact, FSGs are organized in local antennas which belong to a national federation (called FNGN - Federation Nationale des Groupes Naam) in charge of managing the program. The resulting advantages are manifold. First, economies of scale can be reaped through the pooling of food purchases and the collective organisation of transportation from surplus areas. Second, information regarding local food availability and prices easily comes through the federal organizational mode of the network. Monitoring of the use of credit and the management of village granaries is facilitated by the peer pressure arising from continuous comparisons between member units. Credit demands are analyzed during a public meeting organized by the Federation in which village representatives are expected to motivate their credit demand and prove their ability to effectively manage the activity. Credit performance is carefully monitored throughout the year and future access to loans is strictly denied in case of problems. Moreover, when blatant embezzlement occurs, the Federation does not hesitate to pursue the perpetrators in court, thus adding external sanctioning to mutual pressures.

### **3.2 The experimental design**

The program started in 2002 and we took advantage of its scaling-up in 2011 to evaluate its impact on food security. In the area targeted for gradual scaling-up of the program, the NGO had identified eligible villages, that had never benefited from the intervention in the past and had expressed an interest for the intervention. Among these eligible villages, 40 were selected to be part of the experimental framework. Half of them were randomly assigned to the treatment group while the remaining 20

Figure 1: Timing of the intervention and the surveys



villages, used as control units, were to benefit from the program two years later. The intervention consists in setting up a FSG in the village without fixing the level of financial support. While the operational framework is identical in all villages, the amount of credit granted varies across villages and over time, depending on the needs and management capacities of each village.<sup>4</sup>

## 4 Data and descriptive statistics

### 4.1 Data

Our main sample includes 400 households that were surveyed five times during the agricultural years 2010-2011, 2011-2012 and 2012-2013. Figure 1 presents the timing of the intervention and the surveys. The first two surveys were undertaken before and after the 2011 lean season, and include baseline characteristics. The third survey was implemented after the 2012 lean season and coincides with the end of the first year of the intervention. The last two rounds were collected before and after the 2013 lean season. Over this year, the intervention was no longer randomly assigned.<sup>5</sup> As a consequence, our impact assessment relies on the first three rounds of data that is, two rounds before and one round after the treatment.

Based on administrative census, 10 households were randomly selected in each of the 40 villages

<sup>4</sup>The mean credit corresponds to 3,150 euros while all credits granted to the sampled villages were between 1,500 and 5,500 euros.

<sup>5</sup>The renewal of the program revealed to be problematic as the result of embezzlement. A grassroots employee of the Federation who was in charge of 6 villages stole the money entrusted to him to pay back the village loans.

sampled. The sample includes a total of 400 households and 4750 individuals, corresponding to about 5 percent of the population studied. Attrition is low - less than 3 percent of households - and its causes are known and unrelated to treatment assignment.

Broad surveys were implemented in rounds 1 and 4 and more focused follow-up surveys were used in rounds 2, 3 and 5. While general information about the household were asked to the household head, personal information on each adult member and its dependents - e.g. mother and children - were gathered directly from this adult member. Special attention was paid to agricultural production and food stock management, as they are key determinants of food vulnerability. All surveys also include a comprehensive set of questions on food and nutrition. An original section was designed to gather detailed information on all cereal transactions made by household members over the agricultural cycle. On a transaction basis, it includes not only timing, quantities and prices, but also shop characteristics and transaction motives. Also, data on diet diversity, perception of food access and the quality of meals were collected starting from round 3. In addition, we measured and weighted all individuals following WHO standards. We use them to construct indicators for the nutritional status of all households members in each round.

In addition to this main data collection effort, we conducted a detailed investigation on seasonality in the 14 most remote villages. In these villages, a subsample of 70 of the original households was selected to be surveyed on a monthly basis in 2016. Each month, detailed data on food stock management and transactions were collected and all household members were weighted and measured. In the following, we use this data mainly for descriptive purposes.

## 4.2 Descriptive statistics

Tables 1 and 2 provide descriptive statistics. The first table focuses on baseline characteristics and shows that there was no significant difference between treatment and control villages on a large set of village and household characteristics. Table 2 reports food and nutrition indicators for households in control villages and for each of the three agricultural years.

**Nutritional stress** — Panel A of Table 2 reports measures of nutritional status after the lean season and differences in nutritional status before and after the lean season. The measures used include Body Mass Index (BMI) for adults and BMI-For-Age (BFA) z-score for children.<sup>6</sup> These constitute

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<sup>6</sup>Because body fat varies with age and gender during childhood and adolescence, BMI is age and gender specific. Therefore we use a standardized BMI-for-Age z-score, which is defined as the difference between the value for an individual and the median value of a reference (well-nourished) population for the same age and gender, divided by the standard deviation for the reference population. For children below 5, the reference population comes from the WHO

objective measures that are sensitive to short-term variations in food consumption. Based on weight and height, they are proxy measures of adiposity - the amount of fat in the body - and are used as a screening tool to identify individuals who are underweight or suffering from wasting.<sup>7</sup>

We observe that the incidence of malnutrition varies according to age category: 16 percent of adults, 13 percent of children between 5 and 18 years old, 3 percent of children between 4 and 5 years old and 8 percent of children aged 3 or younger were initially identified as underweight. As reflected in changes in both nutritional indices and prevalence rates between 2010-11 and 2011-12, the nutritional situation of all individuals deteriorated over the 2011-12 agricultural year. Children aged 3 or younger were particularly adversely affected: for them, the prevalence of wasting went beyond the 10 percent high level of severity defined by the World Health Organization (WHO, 1995). We also observe that the nutritional status of young children continued to deteriorate in 2012-13. Moreover, variations in the children's BMI between the period preceding and the period following the lean season were quite significant in the years 2010-11 and 2012-13, suggesting a large seasonal stress. This effect is particularly pronounced for the youngest children: the gap between their BMI and that of the well-nourished reference population more than doubled between the two periods.<sup>8</sup> This is an important finding given that seasonal energy stresses are considered as a major contributor to undernutrition. (Vaitla et al., 2009).

The importance of seasonal variations in nutrition is confirmed by the analysis of the monthly data pertaining to the 2016 subsample. Both children above 5 and adults experience a clear decrease in their nutritional level between two harvests (Figures 8 and 10). Interestingly, the drop in adults' BMI coincides with a significant increase in the daily quantity of foodgrain prepared by households (Figure 9). This suggests that the sharp increase in energy expenditure during the period of heavy agricultural work (June to October) is not compensated by the increase in the quantity of food consumed by the household during that period. As children (including young children) participate in agricultural labor, it is not surprising that their nutritional status follows the same trend than that of adults. Our monthly data reveals that very young children (0 to 5 years old) are not protected from noxious

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Child Growth Standard database. It includes a large sample of children from Brazil, Ghana, India, Norway, Oman and United States. The WHO 2007 Growth Reference database provides similar information for children between 6 and 18. We prefer BMI-for-Age to Weight-for-Height because the former can be computed for all children up to 18 while performing equally well in predicting underweight (Mei et al., 2002). Note however that the results presented in the paper hold if we use Weight-for-Height instead.

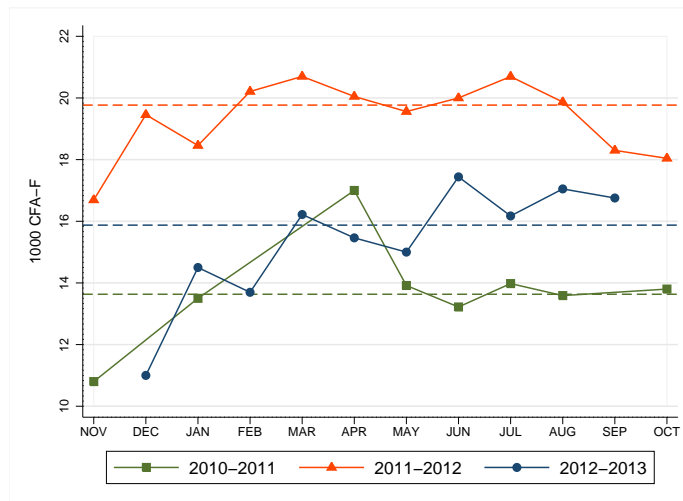
<sup>7</sup>Following WHO (1995), wasting or thinness "indicates in most cases a recent and severe process of weight loss, which is often associated with acute starvation and/or severe disease". According to WHO standards, adults with BMI below 18.5 are underweight. Children and adolescents presenting z-score below -2 suffer from wasting.

<sup>8</sup>Since the BMI-for-age z-score for 0 to 2 years old children is -0.31 after the lean season and since the decrease over the lean season is -0.18, these children went from -0.13 to -0.31 z-score.

fluctuations in nutritional status (Figure 11). Strikingly, this result is driven by children living in the most remote villages: while they experienced a sharp drop in z-score between two harvests in 2016, their counterparts in less remote villages did not.<sup>9</sup> While purely descriptive, this analysis confirms Abey and Hirvonen (2016)’s conclusion regarding the critical role that market access plays to shield the youngest from seasonal food shortages.

**A drought year** — While 65 percent of sampled households produced enough foodgrain to satisfy their needs over the 2010-11 agricultural year, only 13 percent of households were in that situation in 2011-12 (Panel B of Table 2). While there are always some purchases of foodgrain, very high levels are reached after bad harvests. Thus, in 2011-12, purchases amounted to 45 kg per capita, corresponding to about one-third of annual consumption.<sup>10</sup> As illustrated by Figure 2, tight local market conditions translated into very high prices from the very beginning of the agricultural year. The mean price of sorghum was almost 50 percent higher in 2011-2012 than in the previous year, a rate of increase also observed for other crops (FAO et al., 2012). Clearly, the timing of our program evaluation coincides with a drought year, critically raising the potential impact of the intervention.

Figure 2: Monthly Mean Price of Sorghum across Agricultural Year



**Buying further away and earlier** — As evident from panel B of Table 2, almost all cereal transactions take the form of bulk purchases involving 100-kg bags.<sup>11</sup> The main features of these

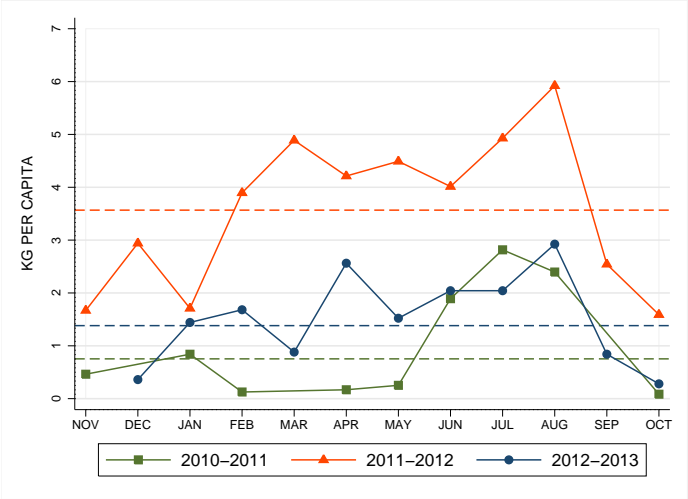
<sup>9</sup>This subsample is drawn from no-road villages. We distinguish between villages where a weekly market takes place and other villages (most remote).

<sup>10</sup>Interestingly, very few households are involved in grain sales while those sales concern negligible quantities. This suggests that households prefer relying on storage rather than on market to smooth consumption within and across years.

<sup>11</sup>This is confirmed when computing the ratio of total cereal bulk purchases to total cereal purchases. More than 99 percent of quantities bought have been acquired through bulk purchases.

transactions are described in panel C. Sorghum is the most important traded foodgrain, far ahead of millet, maize and rice: in 2011-12, it amounted to 65 percent of all grain bought, while maize amounted to 20 percent and millet to 14 percent. While households emphasize their preference for buying close to their dwelling (more on this later), more than half of their purchases are made outside their village. In 2011-12, the situation was even worse since only one-third of the purchased cereals was bought in the village of residence.<sup>12</sup> The timing of purchase is another important dimension to be taken into account. Figure 3 shows that households buy foodgrain through the agricultural year with a small peak during the lean season. In 2011-12, however, a larger proportion of purchased foodgrain (about two-thirds) was acquired before the lean season. This is because stocks started to deplete earlier and households bought larger quantities before own stock depletion.

Figure 3: Monthly Quantity of Grain Bought across Agricultural Year



**Activity of FSGs** — Over the agricultural year 2011-12, each FSG sold an average of 18.1 tons of foodgrain that is, about 3.5 percent of total annual grain requirement.<sup>13</sup> Our data shows that FSG renewed their stock during the year considered (1.7 times on an average), suggesting that a complete temporal arbitrage, from harvest to harvest, did not take place. When analyzing transaction data, we observe that the FSGs’ overall market share was 14 percent while their share rose to 30 percent when only intra-village transactions are taken into account. Almost one-fourth of the households living in

<sup>12</sup>Where existing, traders are typically in small number, they have high cost structures and their activity is of limited duration and scope. Insufficient local food supply exposes populations to adverse market conditions that can eventually end up in food rationing.

<sup>13</sup>This number is calculated using population size and the country consumption reference level of 190 kg of foodgrain per capita per year. When using actual foodgrain consumption, FSGs activities represent 4.5 percent of total annual consumption.

the treatment villages used the local FSG. These two pieces of information indicate that the village granaries are a significant actor in local food markets.

## 5 Methodology and results

In this section, we first investigate the impact of the intervention on food access, the total foodgrain purchased and the timing of purchases. We then explore the effects on final outcomes, namely consumption and nutritional status. We also present heterogeneous effects arising from differences in market integration as measured by the availability of road connections. Because of their isolation, villages where road connections are absent during the lean season (“no-road villages”) are more vulnerable to supply scarcities and therefore prices increase noticeably in times of stress (De Janvry, Fafchamps and Sadoulet, 1991, Newbery, 1989). A total of 16 villages fall into this category and they are equally distributed across treatment and control. Moreover, characteristics of village and households at baseline are well-balanced in the no-road subsample (Table 7).

We mainly use difference-in-difference (DID) estimators and thereby control for time invariant unobservable characteristics. DID allows not only to adjust for initial random differences in mean outcomes across treatment status but also to increase statistical precision, which is important given the limited number of treatment units (Glennerster and Takavarasha, 2013). Specifically, the model we estimate for individual outcomes is:

$$y_{ijt} = \beta_0 + \beta_1 T_j + \beta_2 P_t + \beta_3 T_j P_t + \beta_4' X_{ijt} + \epsilon_{ijt} \quad (8)$$

where  $y_{ijt}$  denotes the outcome of individual  $i$  from village  $j$  at time  $t$ ,  $T_j$  is a binary variable indicating the treatment status of village  $j$ , and  $P_t$  a binary variable taking value 1 for post-intervention observations and value 0 otherwise. The vector  $X_{ijt}$  includes time-varying characteristics such as weather-related indicators and village and year fixed effects. The main coefficient of interest is  $\beta_3$ , which captures the causal effect of the intervention. Because the intervention is implemented at the village level, we systematically cluster standard errors at that level (Bertrand et al., 2004).

### 5.1 Food Access

Table 3 reports the impact of the intervention on foodgrain availability, affordability and purchases. Before turning to our main results, two observations deserve to be made. First, as expected, differences

at baseline across treatment and control units are small and non-significant for all the outcomes considered (first row of Table 3). Second, even after controlling for weather conditions, outcomes in the year of the intervention are significantly different from those at baseline, testifying to substantial pressure on food markets in 2011-12 (second row of Table 3). For example, households had to tread longer distances to acquire food (measured in time spent) and the average price of sorghum was 30 percent higher.

**Foodgrain availability** — The first two columns of Table 3 report the impact of the intervention on local food availability. Column (1) shows that the intervention succeeds in boosting the proportion of foodgrain bought inside the village of residence: the probability that any bag of foodgrain was purchased locally increased by 25.3 percentage points. As shown in Table 8 in Appendix 2, the impact is mainly driven by villages that are not accessible by road, and where availability of foodgrain for local purchase is critically important.

Column (2) reports the impact of the intervention on the total annual distance travelled to buy foodgrain. This aggregate measure is the sum of the amounts of time (in minutes) needed to reach the seller by walk for each transaction. We find that the FSGs allow to significantly reduce the annual distance by an average of 10.3 minutes walk per capita which corresponds to 123 minutes for the average household (a 25 percent reduction of the annual distance travelled by control households). Again the effect is larger in no-road villages (albeit not significantly).

The intervention of the FSGs has clearly succeeded in bringing food closer to rural buyers, which is one of its main goals. When asked to motivate their choice of a particular seller (at baseline), they cited proximity of the seller as the main reason for 78 percent of the quantity of foodgrain purchased (Table 2, panel C). The second most important reason, cited in 20 percent of the cases, is a strong confidence in the actual availability of foodgrain at the selling point. Interestingly lower prices are rarely cited as the main motivation to choose a specific buyer (2 percent). Focus group discussions have highlighted that families prefer to buy foodgrain closer to their dwelling not only because of time and effort gains but also because it reduces the risk of unsuccessful transactions. A transaction is unsuccessful when a villager moves to a nearby market or town to buy foodgrain but returns empty-handed because of unavailability of foodgrain or excessive prices. That the proximity of food shops is a primary concern of the households is evident from their answer to one of our survey questions.

**Foodgrain affordability** — A second important dimension of food security is affordability: food may be available but at such high prices that households cannot acquire it. Column (3) of Table 3



reports the impact of the intervention on nominal sorghum prices paid. In treatment villages, the intervention is responsible for a significant reduction (1,384 CFA-F) of the price of sorghum, which represents a 7 percent cut.<sup>14</sup> It has helped to mitigate the price surge that followed the drought by as much as 25 percent. Again we expect that the price-reducing impact of the intervention will be especially large in remote villages, since remoteness has the effect of isolating a village from price-dampening market forces in times of supply stress. Evidence reported in Table 8 in Appendix 2 confirms this expectation: the price reduction observed in no-road villages is more than four times as high as in the other villages (2445 CFA-F against 586). As seen in rows 3 and 4 in the table, this effect fully annihilates the impact of the drought in remote villages.

**Foodgrain purchase** —So far, we have shown that the FSGs have a positive impact on food access as measured by the availability and affordability criteria: households from treatment villages bought grain closer to their dwellings and at lower prices. What is not clear, however, is the effect of the intervention on the quantities of foodgrain purchased and on the timing of the purchases. We have argued in Section 2 that the former effect is ambiguous in a nutrition-based model with storage losses: it is possible, contrary to immediate intuition, that food purchases do not increase as a result of a decrease in foodgrain prices. As for the latter effect, the model predicts that households will delay their purchases if the intervention assures them that food will be locally available in the lean period.

Columns (4) and (5) of Table 3 provide estimates of the impact of FSGs on the probability for households to have bought any foodgrain and on the annual foodgrain quantities purchased, respectively. First note that throughout the 2011-12 cycle, as many as 80 percent of the households did purchase foodgrain and, on the intensive margin, the quantity purchased per capita was 53 kilograms of foodgrain, that is more than one-fourth of annual requirements.<sup>15</sup> Turning to the impact of the intervention on the two above measures, we find that the parameter estimates are small and not significantly different from zero. In the same line, we observe that the total expenditure on foodgrain has slightly decreased, albeit not significantly (column 6).

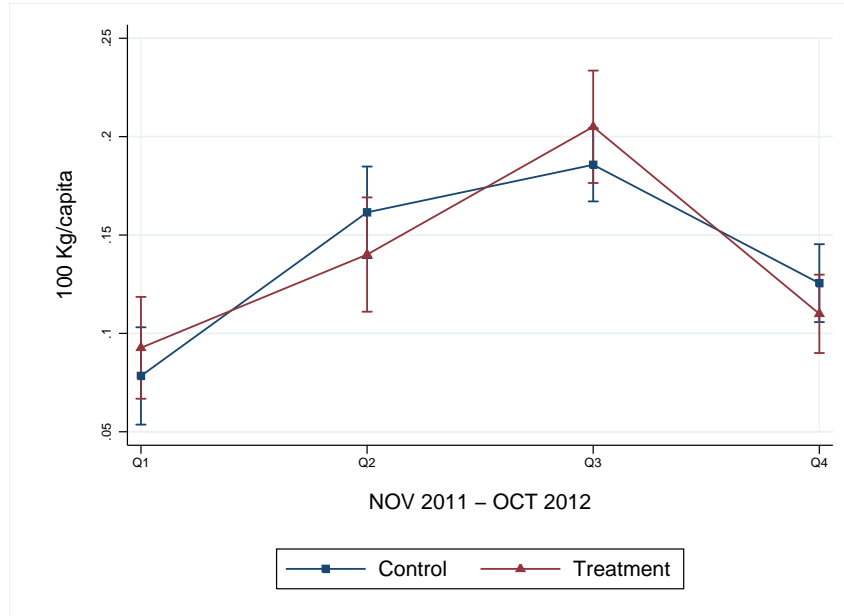
In order to investigate the impact of the intervention on the timing of purchases, we rely on three measures. The first measure consists of the quantities of foodgrain purchased in each quarter of the intervention year. The second one corresponds to the number of months the household holds a stock of foodgrain in the granary, located inside the household compound. The third measure relies on

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<sup>14</sup>Impact on prices reported in Table 3 have been obtained considering each grain bag purchased as the unit of analysis. Results are very similar when considering instead the mean price paid per bag and the household as the unit of analysis.

<sup>15</sup>We use the country consumption reference level of 190 kilograms of foodgrain per capita per year.

Figure 4: FSGs impact on timing of own stock depletion



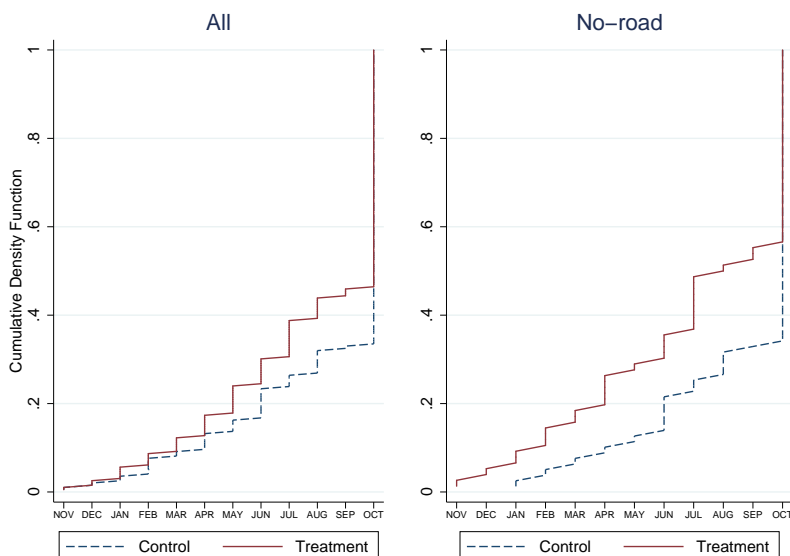
both a binary and a continuous variable that capture purchases made before the depletion of the granary. We know that only grain produced on the family farm goes to the granary, hence the name “own stock” chosen to denote this form of storage. Own foodgrain is stored on ear, while purchased foodgrain is always bought and held in the form of grain inside the household’s main dwelling. A critical observation is that because grain deteriorates faster than ears, the foodgrain purchased is always consumed first, thereby lengthening the duration of the own stock. We actually observe that households do not necessarily deplete their entire granary before starting to purchase foodgrain: a majority of them, 65 percent, extend the duration of their granary stock by purchasing foodgrain before the depletion of their own stock.

Figure 4 compares quarterly purchases in treatment and control villages. It is based on per capita quantities purchased as have been estimated with the help of a negative binomial regression (see regression results in Appendix 2, Table 9).<sup>16</sup> It suggests that households in treatment areas have tended to delay their purchases, compared to control areas. The difference is not statistically significant though.

Turning to the second measure, we expect that households in treatment areas depleted their own stock faster than control households. This expectation is confirmed by Figure 5. The left panel re-

<sup>16</sup>By using a negative binomial regression, we account for the Poisson structure of the quarterly data and the high proportion of zero entries in the data.

Figure 5: FSGs impact on timing of own stock depletion



ports the cumulative distribution of the duration of own stock (in months) for all control and treatment households while the right panel reports the same statistic but only for households living in no-road villages. For both the complete sample and the sample restricted to no-road villages, the cumulative distribution for the control group first-order stochastically dominates the distribution for the treatment group and Kolmogorov-Smirnov tests confirm that the differences across these distributions are statistically significant. We also estimate the impact of the program on the duration of own stock in a regression framework where basic control are introduced. Table 4 reports the results (columns 1 and 2). The intervention shortens the duration of own stock by 0.63 months for the whole sample and by 1.19 months for the restricted sample. Only in the latter case is the effect statistically significant.

Finally, concomitantly to the shortening of the duration of stock, we expect that the intervention reduces anticipated purchases understood as purchases made before the depletion of own stock. Table 4 broadly confirms this prediction. On the extensive margin, households in treatment villages are less likely to make anticipated purchases, but the effect is significant only for no-road villages. In the latter, beneficiaries were 32 percent less likely to make any purchase before stock depletion. On the intensive margin, the intervention lowers substantially the quantities bought before own stock depletion. Indeed, households in treatment villages decrease their anticipated purchases of foodgrain by 11 kg per capita, which represents a 36 percent decrease. It is noticeable that the impact is more than twice as large in no-road villages as in the full sample.

## 5.2 Consumption and Nutrition

We are left with the most important task of examining the impact of the intervention on final outcome variables, food consumption and nutritional status. To assess the impact on nutrition, we rely on anthropometric indicators, as explained earlier. We distinguish between adults (19-59) and three age groups for children (0-2, 3-4 and 5-18). To measure consumption in terms of both quantity and quality, we use a set of household-level variables based on annual grain balance, self-reported daily ration, and recall data on food diet.<sup>17</sup> First, the foodgrain balance, which corresponds to the real foodgrain disposable, is obtained by adding purchases and gifts received to the quantity produced and by then subtracting losses, sales and gifts made.<sup>18</sup> We also compare the real foodgrain disposable to the consumption reference level of 190kg per capita per year and construct a binary variable equal to one if the former quantity exceeds the latter. Second, households have been asked about their daily foodgrain ration during the week preceding the survey (at the end of the lean season), and whether they considered it sufficient to satisfy their basic needs. Third, if foodgrain consumption approximates the caloric intake relatively well in a context where cereals account for two-thirds of this intake, it does not capture the micronutrient adequacy of the diet. To that purpose we construct a diet diversity score (DDS), corresponding to the number of food groups consumed over the last day.<sup>19</sup> As an alternative measure, we use an indicator based on a one-month recall period, which is better adapted to contexts where food diets are very poor, as suggested by Hoddinott (1999). Tables 5 and 6 show the impact of the program on these various outcomes.

**Consumption** — The overall picture that comes out of Table 5 is that there is no clear evidence of an impact of the intervention on food consumption. All coefficients but one are small and not significant. The intervention does not significantly increase the real foodgrain disposable, the probability that the latter exceeds the consumption reference level, and the daily ration. By contrast, treated households are more confident that their daily ration meets their needs. The lack of effect of the FSG program on total foodgrain consumption does not come as a surprise since we earlier found that total foodgrain purchases did not increase in FSG villages.

One could argue that, if the program did not increase the quantity of foodgrain consumed, it may

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<sup>17</sup>As data on food diet were not collected at baseline, impact estimates on this outcome are obtained with simple difference across treatment status.

<sup>18</sup>While different types foodgrain are consumed (mainly sorghum, millet and maize, see above for details), their nutritional content is very similar, both in terms of total energy and micronutrient content. As a result we can sum them up in a unique variable. The results obtained hold if we use other aggregations, based on prices or exact calorie-contents.

<sup>19</sup>Following Steyn et al. (2006), we distinguish between nine foodgroups: (1) cereals, roots and tubers, (2) vitamine-A rich fruits and vegetables, (3) other fruits, (4) other vegetables, (5) legumes and nuts, (6) meat, poultry and fish, (7) fats and oils, (8) dairy, (9) eggs.

have improved the diversity of the diet, as the increase in purchasing power may translate into a greater demand for animal products, vegetables or fruits. That this second effect is also absent is evident from columns (5) and (6) of Table 5: coefficients are small, not significant and even negative. Because the diversity measures only cover the month before the survey, we cannot rule out the possibility that the intervention enabled beneficiary households to improve their diet at other moments of the agricultural cycle (Savy et al., 2006). Nevertheless, evidence from our monthly survey suggests that this is unlikely. While we do observe some changes in food diversity at certain times of the year, they are systematically associated with the availability of some fruits or vegetables that are not purchased.

The analysis of heterogeneous effects across remoteness status leads to a similar conclusion: we do not detect any effect of the FSGs on consumption outcomes.

**Nutrition** — Estimates of the impact on nutritional outcomes are reported in Table 6. The first three columns suggest that the intervention has a large and positive impact on nutritional outcomes for both adults and children. The estimated effect for adults is positive and significant and corresponds to 0.3 BMI point which, on average, corresponds to about one-kilo difference for an individual with mean BMI. Column (2) shows a positive impact on children 5 to 18 years old. Since the post-treatment mean z-scores for this age group in control villages is  $-1$ , the estimated effect of 0.2 z-score in BMI-for-age corresponds to a 20 percent reduction in the existing gap with the well-nourished reference population. Column (4) shows that children below 3 have been positively affected by the program as well. For them the effect corresponds to a 70 percent reduction in the existing gap with the well-nourished reference population. In contrast we do not find any impact of the program for the group of children aged 3 to 4.

Column (5) of Table 6 reveals no significant impact on the prevalence of underweight among adults. More interestingly, the intervention has substantial effect on the incidence of wasting among children. Column (6) thus shows a 4 percentage points decrease in the incidence of wasting among children between 5 and 18, that is, a 25 percent net reduction in the prevalence of wasting. The impact is even greater for the youngest children. In the control villages, the prevalence rate among children under 3 reaches the 10 percent high level of severity defined by the World Health Organisation while it remains stable at moderate levels in the treatment villages. For the sake of comparison, the prevalence of wasting in poor countries is usually below 5 percent in the absence of severe food shortages (WHO, 1995). Again the program appears to have no effect on the nutrition of the group of children aged 3 to 4.

Heterogeneous effects on nutrition are presented in Table 11 in Appendix 2. The effects of the intervention on nutritional outcomes for both adults and children appear mainly driven by the no-road villages.

## 6 Discussion

The presence of FSG enables household to buy foodgrain closer to their dwelling, at lower price and when the need arises rather than in anticipation. The program thus appears to fulfill its objective of activating local food markets. The ultimate goal of improving nutrition is reached as well: FSG enhance children's as well as adults' nutritional status. Nevertheless, the lack of impact on total quantity of food consumed deserves elucidation. Two questions arise. First, why is it that households have not increased the quantity of grain purchased while prices have decreased? And second, how can we account for an improvement in nutrition when total grain available (the sum of own production and purchase) has not increased as a result of the program?

### **Ruling out Giffen and quality effects**

The Giffen effect constitutes the most straightforward explanation for both facts. The decrease in foodgrain price leads to an increase in purchasing power that induces the households to diversify their food diet away from foodgrain. If this income effect outweighs the substitution effect, we expect a net decrease in foodgrain consumption. Recent evidence from China confirms that the Giffen effect may be observed in contexts that resemble ours: households are poor and obtain most of their calories from the consumption of staple grains (Jensen and Miller, 2008). The second fact can also be explained by the Giffen effect if an increase in food diversity improves the quality of nutrition (Steyn et al., 2006).

Surprisingly however, our evidence does not support this explanation. As seen in Table 5, there is no impact of the program on various food diversity scores (at least at the end of the lean season). Of course, diversification needs not concern only food: an increase in real income may prompt households to increase the consumption of other goods and services that have a positive influence on nutrition. In particular, health expenditures could increase and improve nutrition to the extent that healthier individuals have a more efficient metabolism and better absorb nutrients. We actually have a detailed measure of health expenditures yet, unfortunately not for the year of the intervention: health expenditures have been measured only at the baseline and two years after the start of program. If we cannot rule out an effect of the program during the year of the intervention, qualitative evidence runs against

this interpretation. The sample households, indeed, confessed to not using preventive medicine and to have recourse to medical treatment (conventional or traditional) only as a last resort solution. The data confirm that health expenditures are small (2 percent of total cash expenditure). Furthermore, we find no impact of the program on the occurrence and duration of episodes of disease for children and adults, suggesting that the improvement of nutritional outcomes does not result from a reduction in disease exposure in treatment villages.<sup>20</sup>

Different from a Giffen effect is an explanation based on a change in the quality of foodgrain. Thanks to a higher nutrition content of a given quantity of cereal, households would be able to improve their nutrition status as a result of the program, even if they do not increase the quantity purchased. Again, our evidence does not support this second explanation. First, a change in the quality of cereals was never mentioned by the sample households when we asked them about the advantages of the program (in an open question). Second, if this explanation was relevant, we would expect no impact of the program for households who did not purchase cereals in the FSGs. Table 12 in Appendix 2 indicates that this is not the case.

Note finally that the impact of the program on nutrition does not seem to be driven by a reduction in energy expenditures. First the reduction in the travel distance to acquire cereals is too small to explain any significant increase in weight among households in FSG villages (it represents less than 1000 kcal per household per year). Furthermore, we find no evidence that households in FSG villages have exerted less effort as reflected in the activities undertaken or in the amount of agricultural production in the post-intervention campaign.<sup>21</sup> If anything yields have slightly improved.

### **Intertemporal reallocation of purchases and stock management imperfections: quantitative evidence**

In line with our model (see Section 2), Table 4 has shown that the program significantly affects the timing of food purchases. Specifically, households in control areas tend to anticipate scarcity by buying food even though they still have foodgrain available from their own harvest. By contrast, being more assured of future supply, households in treatment villages tend to postpone their purchases until their own stock is depleted. At any point of time, therefore, they have smaller quantities of foodgrain stored in their household. In the presence of storage losses, they may thus enjoy higher effective consumption

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<sup>20</sup>Results available upon request.

<sup>21</sup>There is no effect of the program neither on the propensity of treated individuals to engage in income generating activities nor on the income generated by these activities (see Table 13 in Appendix 2).

than households in control villages, even though they purchase the same quantity of foodgrain.<sup>22</sup>

We have already mentioned two forms of physical storage: storage of own production in the household granary in the form of ears, and storage of purchased food inside the dwelling in the form of grain. There also exists the possibility of storing in the body, whereby body mass is accumulated and then gradually depleted during the lean season when consumption of cereal decreases and/or effort increases. There is a literature investigating this mechanism in the context of subsistence farming (Dugdale and Payne, 1987, Branca et al., 1993, IFPRI, 2015).

All these forms of storage obviously entail costs. For body mass storing, there are costs associated with storing and de-storing as well as costs associated with the maintenance of a larger body mass (Dugdale and Payne, 1987).<sup>23</sup> For foodgrain storing, two different types of costs are involved. The first and most well-known type consists of foodgrain losses caused by rodents and moisture. The second type results from the difficulty to protect household grain from demands of visitors and household members themselves. When the stocks of food are ill-protected against non-household members, the source of storage costs lies in redistributive pressures (Platteau, 2000, Baland et al., 2011, Dupas and Robinson, 2013, Jakiela and Ozier, 2015). When they are ill-protected against household members themselves, the source lies more in an inability to withstand the pressure to consume purchased foodgrain quickly. This inability is akin to a self-control problem (Ashraf et al., 2006, Bernheim et al., 2015). Regarding redistributive pressures, interviews with sample farmers reveal that large household stocks signal abundance and attract solicitations. In particular, visitors are likely to stay longer. Regarding the urge to consume, the idea is that people may find it difficult not to consume food that is readily accessible and in apparent abundance. The problem is expected to be especially acute when people are hungry. Finally, it may be noted that body mass accumulation may itself be the consequence of a self-control problem, in this case the urge to eat when food is plentiful.

What evidence do we have to support the story based on storage cost? First, physical losses turn out to be much less important than we expected. Only 1.5 percent of households in the sample declared that they had suffered any loss due to physical storage problems, and the quantities concerned are always small (never more than 5 percent). This result confirms the finding of a recent World Bank Report that storage losses are small in dry and semi-dry areas (World Bank, 2011). Second, turning to

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<sup>22</sup>It is interesting to note that storage costs play a role similar to improved quality of foodgrain purchased.

<sup>23</sup>A back-of-the-envelope calibration suggests that the magnitude of the impact on nutritional status after the lean season is easily compatible with a more efficient timing of consumption, unaccompanied by an increase in the total quantity consumed. Thus, one additional kilogram gained before the lean season is completely lost after a period of 5 months if no compensatory energy is consumed in the meantime for its maintenance (for a moderately active woman, FAO, 2001). By smoothing weight over this period, such a loss can be avoided and a net gain can be obtained.



evidence about redistributive pressures, we note that while treatment households received fewer visits of people staying and eating in the household, the effect is not quantitatively significant. In fact, in the analysis presented above, we have accounted for these visits when computing the quantity of grain consumed by household members, and, as we know, there is no significant effect of the program on the quantity of food consumed per capita (see Table 5). Redistribution can also take the form of gifts of cereals to non-residents. We have detailed information on such transfers made over the agricultural year. It appears that there is a significant difference in such transfers between households in the treatment and control villages. However, the quantities concerned are small (see Table 13 in Appendix 2) and they have been taken into account when computing the total foodgrain disposable. We must therefore conclude that even though there may be a mitigating impact of FSGs on redistributive pressure, it is too insignificant to explain our results.

Third, the existence of a self-control problem as defined above cannot be formally tested on the basis of our data. Still we find supportive evidence, both quantitative and qualitative, that cannot be easily dismissed. Quantitatively, we would have liked to measure the impact of the program on the intra-year fluctuations in body mass in order to test whether these fluctuations have been dampened. Unfortunately, because we have only one wave of survey during the year following the intervention, we are unable to carry out this test. We can nevertheless adduce indirect evidence based on data on body mass fluctuations observed during years 2010-11 and 2012-13. It shows that the delaying of foodgrain purchases is associated with less body mass fluctuations. Specifically, the variation in adult body mass for households who purchased cereals after depletion of own stock is significantly smaller than the variation for households who made anticipated purchases. The same significant difference is observed when, instead of comparing households which did or did not make anticipated purchases, we use a continuous variable consisting of the quantities purchased before stock depletion. Table 14 in Appendix 2 indicates that anticipated purchases (made before stock depletion) are associated with higher body mass indices before the lean season (column 1 or 4) but similar body mass indexes after the lean season (columns 2 or 5), implying a higher variation in body mass compared to the other households (columns 3 or 6). These findings suggest that because it induced households to limit their anticipated purchases, the program also led to a reduction in body mass fluctuations. Better timing of cereal purchases thus appears as an effective way to dampen the urge to quickly consume readily accessible foodgrain. This conclusion is supported by the analysis of the relationship between the timing of purchase and the quantity of food prepared in households in the subsample of households

surveyed monthly in 2016. Controlling for the annual foodgrain disposable, households appear to prepare significantly more food right after having made a purchase (Table 15 in Appendix 2). Thus delaying purchase until the need arises in the lean season may help smooth nutrition.

### **Intertemporal reallocation of purchases and stock management imperfections: qualitative evidence**

In order to ascertain our interpretation of the results we have returned to the field to conduct in-depth interviews in the framework of follow-up workshops organized in both treatment and control villages (June 2015). In a first step, we devised visual tools for the purpose of summarizing our most salient quantitative findings in an easily understandable manner (see Figures ?? to Figure 15 in Appendix 3). Group discussion was then aimed at eliciting opinions about these findings and their interpretation. In particular, participants were explicitly asked whether the paradox discovered - quantities of foodgrain consumed have not been affected by the intervention yet the nutritional status has improved - was an artefact born of ill-measured variables and, if not, what could possibly explain it.

In a second step, we used boards that allowed individual participants to illustrate their stock management and consumption strategies (see photo in Figure 18).<sup>24</sup> Specifically they were given twelve cards representing the monthly rations available for their household: eight of them were quantities drawn from their own stock and the four remaining cards corresponded to purchases. They were asked to allocate these cards month by month so as to allow us to visualize the timing of their purchases. Participants were then asked to justify their choice. A striking outcome of this exercise was the emergence of two neatly differentiated time patterns: one in which purchases occurred rather early, that is, before the lean season, and the other in which they occurred later (Figure 16 of Appendix 3). Local availability of foodgrain during the lean (rainy) season came out as the most important concern guiding their choice. Subsequently, in the light of their purchase pattern, participants were asked to indicate month by month the daily quantities of foodgrain prepared by their household. Their choice was restricted to three possibilities: a big, a medium and a small bowl. The main lesson here is that households who purchased earlier also tended to consume greater quantities during the months of purchase. Figure 17 of Appendix 3 illustrates two canonical patterns. In the left panel, the household purchases early and consumes relatively large quantities of food before the lean season. In the right panel, by contrast, purchases are delayed and consumption improves later in the year when agricultural

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<sup>24</sup>A total of 15 individuals participated in this activity.

work is at its highest. This exercise highlighted the existence of a pressure-to-consume problem that has been confirmed in in-depth individual interviews conducted afterwards. Interviewees, indeed, recurrently mentioned and documented how the temptation to quickly consume foodgrain within easy reach drive their consumption time pattern. Such temptation appears to be especially strong among mothers who cannot bear the sight of their hungry children: “we are the ones who have to calm down the children when they cry of hunger during the night”, said one of the interviewed women. Revealingly, household heads admitted that it is hard for them to resist the pressure of their wife (wives), particularly when bags of purchased foodgrain are available inside the dwelling.

Recent technical reports from various inventory credit (warrantage) programs implemented in Burkina Faso point to the same interpretation. Inventory credit programs provide credit against the deposit of cereals in community granaries. The purpose of warrantage is to relax the farmers’ liquidity constraints and allow them to avoid the costly “sell low, buy high” behavior - producers sell foodgrain at low price just after the harvest and buy it back at high price during the lean season. A striking lesson from these reports, however, is that pressure-to-consume and redistributive pressure are major issues confronting the households in managing their cereal stocks. Thus Ghione et al. (2013) note that during the 2012-13 campaign 17 percent of bags stored belong to producers who did not request a loan yet paid for the storage. To explain this counter-intuitive behavior the authors mention two effects. The ability to store food outside of the compound enables the household not only to reduce the quantity of food consumed by the family itself but also to reduce the food distributed to other members of the community as a result of social pressure. In the words of a program beneficiary, since home storage attracts repeated demands from family members, “storing at home entails losses, and the family is the most damaging pest”. The report by Oxfam International (2015) goes into the same direction: having less foodgrain readily accessible inside the compound has the advantage of mitigating social pressure, itself justified in terms of solidarity obligations. Moreover, the households are protected against the temptation to sell grain as soon as need arises. Hence households reach the lean period with greater quantities of foodgrain than in the absence of the warrantage program. For Coulter (2014), finally, households view warrantage as a form of forced savings and as a way to withdraw part of their harvest from the sight of their close kin or to avoid the temptation to sell cereals to finance weddings, baptisms, funerals, etc ... A new insight that emerges from the above statements is that social pressure and pressure to consume are intimately related. The pressure arises not only vis-a-vis the drive to consume today what is better left for tomorrow but also vis-a-vis the drive to satisfy social needs,

including helping relatives or villagers.

Let us now sum up our story. As a result of the program households feel more secure in their access to foodgrain: they believe that foodgrain will be readily available throughout the year, at reasonable prices, and within rather short distances. We know that this perception aptly reflects reality in treatment villages. To describe their feeling of security, people use a colourful expression: the program has brought them “the peace of the heart” (la paix du coeur). Feeling less anxious about future availability of foodgrain, they are more willing to purchase cereals as the need arises. In other words, they refrain from anticipated purchases and thus avoid the costs of storage, direct or indirect. In particular, they may reduce body fat accumulation which is a second-best strategy in a context of food shortage.

## 7 Conclusion

This paper makes three important contributions. First, it confirms that, especially in remote areas where local markets are thin, food market activation has the effect of smoothing interseasonal nutritional status. The effect is strongest among children, and young children in particular, for whom deficient nutrition has devastating long-term consequences. Second, and surprisingly, this beneficial effect is obtained despite the fact that total food consumption has not increased as a result of the external intervention. With the help of a simple two-period model, we show that an increase in consumption needs not take place when the price of foodgrain declines during the lean season and the household optimally adjusts its consumption behavior to the change in price. The question then arises as to how nutritional status can improve in the absence of an increase in consumption. The answer to this question constitutes our third key finding: a change in the timing of food purchases translates into a change in the timing of consumption that drives the nutritional improvement. The underlying mechanism is the better ability of the household to mitigate food storage imperfections understood in a broad sense. Being assured of a more reliable supply of foodgrain in the lean season, households choose to first consume their own stock before starting to purchase foodgrain. In other words, they postpone their purchases, which allows them to economize on the costs of storage. More than the waste of the foodgrain stored, these costs mainly consist of an ineffective distribution of consumption over time due to excessive consumption of foodgrain purchased before the lean season (before the stocks are depleted). The problem is one of pressure-to-consume that is aggravated by the fact that, unlike the harvest grain stored on ears in the household granary, food purchased is kept in the house in an

immediately accessible and eatable form.

The problem of storage imperfections as understood above has not received adequate attention in the literature dealing with nutritional stress and savings behavior. This paper has offered a first and necessarily incomplete approach towards explaining the behavior of households subject to nutritional stress in conditions of highly imperfect foodgrain markets. The important role of losses stemming from sub-optimal timing of food consumption is an unexpected finding of our empirical study. This explains why our investigation tools were not designed to address this issue systematically, in particular to formally test for the presence of a self-control problem. We leave this task for future research.

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Figure 6: Nutrition in each period as a function of  $P_2$  ( $P_1 = 1$ )

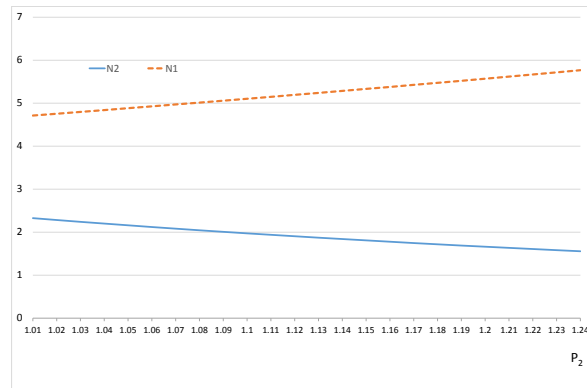


Figure 7: Total quantity consumed over the two periods as a function of  $P_2$  ( $P_1 = 1$ )

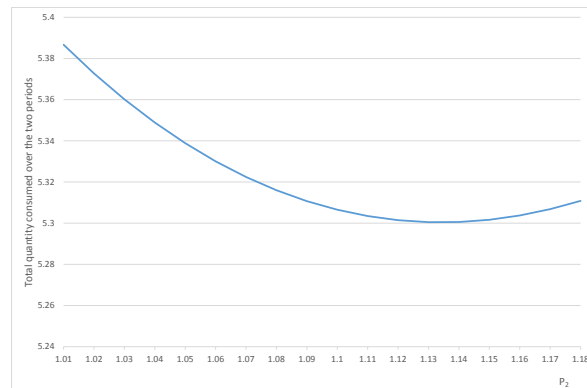


Figure 8: Seasonal variations in adults BMI (2016 subsample)

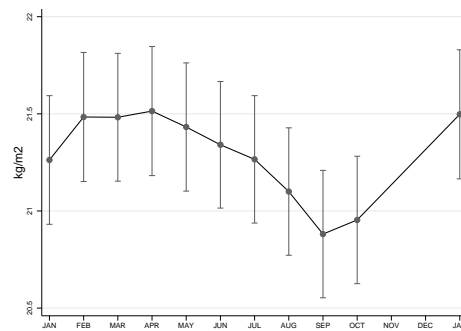


Figure 9: Seasonal variations in daily per capita foodgrain ration (2016 subsample)

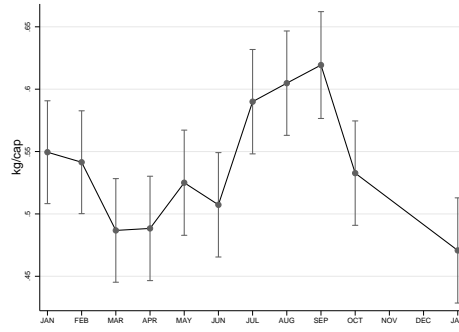


Figure 10: Seasonal variations in BMI-for-age for children age 5 to18 (2016 subsample)

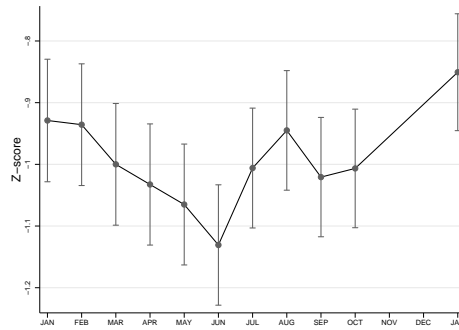


Figure 11: Seasonal variations in BMI-for-age for children below 5, by remoteness (2016 subsample)

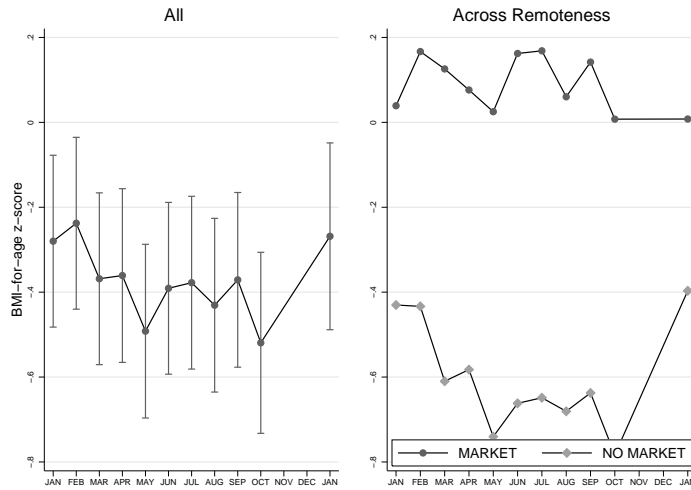


Table 1: Descriptive statistics and balance tests on baseline characteristics

	Treatment (T)			Control (C)			(T) - (C)		
	N	MEAN	SD	N	MEAN	SD	DIFF	SE	
<b>Village-level characteristics</b>									
Village population (# individuals)	20	2735.20	3018.77	20	3793.55(4)	6619.97	-1058.35	1626.91	
Distance to the nearest community health center (km)	20	2.35	3.75	20	3.25	3.60	-0.90	1.16	
Distance to the nearest town (km)	20	17.00	7.58	20	15.35	9.68	1.65	2.75	
=1 if no road passing through the village	20	0.50	0.40	20	0.40	0.50	-0.40	0.16	
Distance to the nearest road (km)	20	3.15	4.76	20	4.25	6.53	-1.10	1.81	
=1 if no market place in the village	20	0.50	0.51	20	0.50	0.51	0.00	0.16	
Distance to the nearest market place (km)	20	3.50	4.14	20	3.35	4.70	0.15	1.40	
=1 if no permanent cereal trader in the village	20	0.70	0.47	20	0.65	0.49	0.05	0.15	
Transport cost city-village (in CFA-F/sack of grain)	20	642.50	408.23	20	655.00	377.28	-12.50	124.30	
End-of-season harvest indicators (2011 WRSI for sorghum (5))	20	84.35	9.21	20	87.15	11.12	-2.80	3.23	
=1 if 2011 rain started late (in July)	20	0.55	0.51	20	0.50	0.51	0.05	0.16	
=1 if 2011 precipitations were less abundant than usual	20	0.95	0.22	20	1.00	0.00	-0.05	0.05	
<b>Household-level characteristics</b>									
Household (HH) size (# HH members)	200	11.98	5.36	200	11.94	5.92	0.05	0.74	
Number of HH members below 14	200	6.18	3.17	200	6.18	3.92	0.00	0.50	
=1 if polygamous HH	200	0.62	0.49	200	0.56	0.50	0.06	0.07	
=1 if male household-head (HH-H)	200	0.98	0.12	200	0.98	0.12	0.00	0.01	
Age of HH-H	200	54.73	13.84	200	54.51	14.34	0.21	1.91	
=1 if HH-H native from village	200	0.95	0.22	200	0.92	0.28	0.03	0.03	
=1 if HH-H Mossi (main ethnic group)	200	0.90	0.30	200	0.74	0.44	0.15	0.10	
=1 if HH-H muslim (main religious group)	200	0.81	0.40	200	0.79	0.41	0.02	0.08	
=1 if HH-H close relative of a village leader	200	0.47	0.50	200	0.43	0.50	0.04	0.05	
=1 if HH-H went to formal school	200	0.36	0.48	200	0.38	0.49	-0.01	0.06	
=1 if HH-H part of a village organisation	200	0.20	0.40	200	0.20	0.40	0.00	0.05	
=1 if house made of concrete wall	200	0.05	0.22	200	0.04	0.20	0.01	0.02	
=1 if HH owns a motorcycle	200	0.38	0.49	200	0.45	0.50	-0.07	0.05	
=1 if any small business	200	0.54	0.50	200	0.63	0.48	-0.09	0.06	
=1 if HH owns some livestock	200	0.97	0.16	200	0.95	0.22	0.02	0.02	
Cattle size (# of head)	200	20.27	21.14	200	18.63	18.98	1.64	2.32	
Surface of land cultivated (Ha/cap)	199	0.28	0.16	192	0.29	0.16	-0.00	0.02	
=1 if self-sufficient in cereals over the last 3 years	199	0.39	0.49	198	0.38	0.49	0.00	0.07	
2011 cereal production (kg/cap)	196	107.51	118.51	194	107.49	96.11	0.02	15.64	
PPI consumption index (6)	200	20.43	6.12	200	21.54	7.25	-1.10	0.94	
Annual total expenditures (in 1000 CFA-F/cap)	200	73.84	38.11	200	81.51	39.37	-7.67	4.69	
Share of food expenditures	200	0.73	0.14	200	0.72	0.16	0.01	0.02	
Share of health expenditures	200	0.02	0.03	200	0.02	0.07	-0.00	0.01	

(1) With exceptions (1 household was not involved in any grain production), missing values are due either to the absence of the respondent or to unavailable information.

(2) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

(3) Standard errors (SE) corresponds to village-level Cluster-Robust-Standard-Errors (CRSE).

(4) Higher population size in controls is explained by the presence of a small city in this subsample. Otherwise, village sizes are about the same in the two groups - on average 2500 inhabitants.

(5) WRSI is a water balance model that is used by Food and Agricultural Organization (FAO) and FEWS NET scientists to provide crop yield assessments (for more details, see Verdin, 2002).

(6) The Progress out of Poverty Index for Burkina Faso is a poverty measurement tool based on eight low-cost indicators to estimate the likelihood that a household has consumption below a given poverty line (for more details, see Schreiner, 2011).

Table 2: Nutritional outcomes and foodgrain consumption across agricultural cycles

	2010-2011		2011-2012		2012-2013	
	MEAN	SD	MEAN	SD	MEAN	SD
<b>Panel A : Nutritional Outcomes (Individual level)</b>						
BMI level (after lean season) of adults (19-59 years old)	20.79	2.52	20.57	2.56	20.68	2.56
BMI-for-age z-score of 6-18 year old children	-0.96	0.97	-1.00	0.99	-1.05	1.07
BMI-for-age z-score of 4-5 year old children	-0.23	1.02	-0.27	1.06	-0.39	1.06
BMI-for-age z-score of 0.5-3 year old children	-0.31	1.33	-0.45	1.27	-0.49	1.63
=1 if adults BMI <18.5	0.16	0.37	0.18	0.39	0.18	0.39
=1 if 6-18 year old children BMI-for-age <-2	0.13	0.34	0.15	0.36	0.18	0.38
=1 if 4-5 year old children BMI-for-age <-2	0.03	0.17	0.05	0.22	0.07	0.26
=1 if 0.5-3 year old children BMI-for-age <-2	0.08	0.27	0.12	0.33	0.17	0.37
Difference (after-before lean season) in BMI level of adults	0.04	1.41	-	-	-0.74	1.26
Difference in BMI-for-age of 5-18 year old children	-1.04	1.14	-	-	-0.79	0.99
Difference in BMI-for-age of 4-5 year old children	0.03	1.17	-	-	-0.16	1.21
Difference in BMI-for-age of 0.5-3 year old children	-0.18	1.54	-	-	-0.36	1.32
<b>Panel B : Foodgrain Production and Market Participation (Household level)</b>						
Foodgrain production (Kg/cap)	242.47	145.38	104.31	101.47	158.27	110.45
=1 if foodgrain self-sufficient	0.65	0.48	0.13	0.34	0.43	0.50
=1 if any foodgrain sale	0.02	0.14	0.04	0.20	0.07	0.25
Foodgrain sales (Kg/cap)	0.70	5.96	0.61	3.23	2.04	9.51
=1 if any foodgrain purchase	0.51	0.50	0.83	0.38	0.42	0.50
Foodgrain purchases (Kg/cap)	-	-	45.23	42.93	17.89	33.95
=1 if any foodgrain bulk (>100 kg) purchase	0.34	0.47	0.76	0.43	0.38	0.49
Foodgrain bulk (>100 kg) purchases (Kg/cap)	10.50	20.59	45.15	42.94	17.50	34.04
<b>Panel C : Foodgrain Purchases (Transaction level)</b>						
=1 if sorghum	0.72	0.45	0.64	0.48	0.47	0.50
Nominal price paid for sorghum	13.64	2.24	19.77	3.36	15.87	2.66
=1 if bought in the village	0.50	0.50	0.37	0.48	0.45	0.50
Nominal price paid for sorghum in the village	13.82	1.55	20.39	3.01	17.11	2.32
=1 if bought during lean season	0.79	0.41	0.38	0.49	0.45	0.50
Nominal price paid for sorghum during lean season	13.66	2.28	20.17	3.75	16.90	2.41
=1 if bought before own stock depletion	0.43	0.50	0.65	0.48	0.57	0.49
=1 if bought to a particular seller because of - Proximity	0.78	0.42	0.39	0.49	0.43	0.50
=1 if bought to a particular seller because of - Availability	0.20	0.40	0.36	0.48	0.47	0.50
=1 if bought to a particular seller because of - Price	0.02	0.14	0.18	0.39	0.09	0.29

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- (1) All figures are drawn from the sample of 200 households from contol villages.
- (2) Foodgrain retail purchases were not investigated in depth at baseline. This explains why bulk purchases correspond to total purchases in 2010-2011.
- (3) Panel C includes all bulk purchases of foodgrain and the unit of observation is the 100-kg bag of foodgrain. There are 439 bags of foodgrain bought in 2010-2011, 2076 in 2011-2012 and 670 in 2012-2013.

Table 3: Impact of FSGs on foodgrain access

	AVAILABILITY		AFFORDABILITY		PURCHASE	
	=1 IF GRAIN BAG BOUGHT LOCALLY	ANNUAL DIST. TRAVELLED PER CAPITA IN MINUTES	PRICE PAID BAG OF SORGHUM IN 1000 CFA-F	=1 IF HH MAKES ANY GRAIN PURCHASE	ANNUAL QUANTITY BOUGHT IN 100KG/CAP <sup>(7)</sup>	ANNUAL EXPENDITURES PER CAPITA IN 1000 CFA-F
	(1)	(2)	(3)	(4)	(5)	(6)
DIFF. AT BASELINE <sup>(5)</sup>	-0.130 [ 0.177]	1.038 [ 2.103]	0.403 [ 0.502]	-0.015 [ 0.056]	-0.004 [ 0.025]	-0.092 [ 0.342]
AFTER (=1 IF POST-TREAT)	0.027 [ 0.114]	37.971*** [ 6.081]	5.767*** [ 0.607]	0.521*** [ 0.054]	0.443*** [ 0.064]	7.366*** [ 1.322]
AFTER * TREAT (DID)	0.253* [ 0.143]	-10.276* [ 5.920]	-1.384** [ 0.593]	0.030 [ 0.056]	-0.034 [ 0.057]	-0.703 [ 1.136]
CONTROL GROUP MEAN <sup>(6)</sup>	0.367	44.127	19.769	0.817	0.532	8.591
LEVEL OF ANALYSIS OBSERVATION	BG 2516	HH 791	BG 1628	HH 791	HH 791	HH 791

- (1) The dependent variables are at bag of grain (BG) or household (HH) level.  
(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogeneous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(3) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).  
(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .  
(5) Baseline difference in outcomes are computed through separate OLS regressions.  
(6) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.  
(7) Annual foodgrain quantity bought includes retail transactions in addition to bulk purchases.

Table 4: Impact of FSGs on anticipated purchases

	NUMBER OF MONTHS BEFORE OWN STOCK DEPLETION <sup>(3)</sup>		=1 IF ANY CEREAL PURCHASE BEFORE OWN STOCK DEPLETION		QUANTITY PURCHASED BEFORE OWN STOCK DEPLETION (100 KG/CAP)	
	ALL	NO ROAD	ALL	NO ROAD	ALL	NO ROAD
	(1)	(2)	(3)	(4)	(5)	(6)
TREAT	-0.631 [ 0.422]	-1.189* [ 0.665]	-0.077 [ 0.054]	-0.302*** [ 0.053]	-0.110** [ 0.050]	-0.260*** [ 0.082]
CONTROL GROUP MEAN <sup>(7)</sup>	10.325	10.506	0.650	0.696	0.308	0.340
LEVEL OF ANALYSIS OBSERVATION	HH 393	HH 155	HH 393	HH 155	HH 393	HH 155

- (1) Estimates correspond to simple difference across treatment status and not to differences-in-differences using baseline.  
(2) In columns (1) and (2), the dependent variable corresponds to the number of month since harvest (october 2011).  
(3) The own stock refers to the stock of foodgrain produced on farm.  
(4) All regressions include control variables corresponding to village-level time-varying exogeneous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(5) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).  
(6) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .  
(7) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

Table 5: Impact of FSGs on food consumption

	DISPOSABLE		RATION		DIVERSITY	
	LN OF REAL GRAIN DISPOSABLE IN KG/YEAR/CAP	=1 IF REAL GRAIN DISPOSABLE > STANDARD	SELF-REPORTED GRAIN DAILY RATION IN KG/DAY/CAP	=1 IF DAILY RATION CONSIDERED AS SUFFICIENT	HODDINOTT DIETARY DIVERSITY SCORE	IFPRI DIETARY DIVERSITY SCORE
	(1)	(2)	(3)	(4)	(5)	(6)
DIFF. AT BASELINE <sup>(8)</sup>	0.086 [ 0.080 ]	0.045 [ 0.065 ]	-0.003 [ 0.007 ]	-0.055* [ 0.029 ]	-	-
AFTER (=1 IF POST-TREAT)	-0.468*** [ 0.101 ]	-0.406*** [ 0.076 ]	0.044*** [ 0.008 ]	-0.108** [ 0.051 ]	-	-
AFTER * TREAT (DID)	-0.129 [ 0.089 ]	-0.080 [ 0.065 ]	-0.011 [ 0.008 ]	0.086** [ 0.039 ]	-11.934 [ 7.946 ]	-0.035 [ 0.065 ]
CONTROL GROUP MEAN <sup>(9)</sup>	4.930	0.281	0.183	0.845	187.345	3.964
LEVEL OF ANALYSIS OBSERVATION	HH 780	HH 780	HH 786	HH 786	HH 393	HH 393

(1) Real grain disposable corresponds to (production + purchases + gifts in - losses - sales - gifts out).

(2) The consumption standard in Burkina Faso corresponds to 190 kg/year/capita or, equivalently, 0.520 kg/day/cap.

(3) Hoddinott (2001) dietary diversity score takes into account all food items as such and the frequency of their consumption (3-level index) over the last month. IFPRI Dietary Diversity Score relies on 9 food categories for its construction and corresponds to the number of food categories consumed over the last day. Here, it has been computed as the number of food categories consumed over the last month.

(4) All regressions include village fixed effects (except regressions from columns 5 and 6) and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.

(5) Estimations in column (5) and (6) are simple differences. Baseline data on food diet do not exist.

(6) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).

(7) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

(8) Baseline difference in outcomes are computed through separate OLS regressions.

(9) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

Table 6: Impact of FSGs on nutrition

	LEVEL				PREVALENCE			
	BMI	BMI-F-AGE Z-SCORE			=1 IF BMI <18.5	=1 IF BMI-F-AGE Z <-2		
	19-59	5-18	3-4	0-2 <sup>(1)</sup>	19-59	5-18	3-4	0-2
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
DIFF. AT BASELINE <sup>(5)</sup>	-0.069 [ 0.228 ]	-0.095 [ 0.077 ]	0.037 [ 0.137 ]	-0.177 [ 0.136 ]	0.006 [ 0.031 ]	0.005 [ 0.022 ]	0.007 [ 0.016 ]	0.037** [ 0.018 ]
AFTER (=1 IF POST-TREAT)	-0.075 [ 0.108 ]	-0.072 [ 0.044 ]	0.050 [ 0.128 ]	-0.373** [ 0.147 ]	0.009 [ 0.021 ]	0.024 [ 0.019 ]	0.005 [ 0.031 ]	0.060* [ 0.031 ]
AFTER * TREAT (DID)	0.301** [ 0.140 ]	0.197** [ 0.052 ]	-0.017 [ 0.131 ]	0.366** [ 0.165 ]	-0.022 [ 0.022 ]	-0.040* [ 0.023 ]	-0.018 [ 0.030 ]	-0.088** [ 0.041 ]
CONTROL GROUP MEAN <sup>(6)</sup>	20.575	-1.012	-0.261	-0.513	0.183	0.148	0.044	0.115
LEVEL OF ANALYSIS OBSERVATION	IND 2329	IND 3069	IND 623	IND 721	IND 2329	IND 3069	IND 623	IND 721

(1) The 0-2 year age category corresponds to children from 6 to 36 months.

(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.

(3) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).

(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

(5) Baseline difference in outcomes are computed through separate OLS regressions.

(6) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

## Appendix 1

In this appendix we derive the condition for a positive purchase in period 2 and expressions (6) and (7). We also introduce the functional forms used for the simulation.

### Lagrangian and first-order conditions

The maximization problem yields the following Lagrangian, where non-negativity constraints for  $m_1, m_2$  are included and  $\lambda_1, \lambda_2, \lambda_3, \lambda_4, \nu_1, \nu_2$  are (non-negative) Lagrangian multipliers:

$$\begin{aligned} L = & U(N_1, O_1) + \gamma U(N_2, O_2) - \lambda_1 (N_1 - f(N_0) - n(C_1) + a(L_1)) - \lambda_2 (N_2 - f(N_1) - n(C_2) + a(L_2)) \\ & - \lambda_3 (rP_1m_1 + rO_1 + P_2m_2 + O_2 - rR) - \lambda_4 (\beta C_1 + C_2 - \beta m_1 - m_2) \\ & - \lambda_5 (C_1 - m_1) + \nu_1 m_1 + \nu_2 m_2 \end{aligned}$$

The first order conditions are:

$$\frac{dL}{dO_1} = U_O(N_1, O_1) - \lambda_3 r = 0 \quad (9)$$

$$\frac{dL}{dO_2} = \gamma U_O(N_2, O_2) - \lambda_3 = 0 \quad (10)$$

$$\frac{dL}{dN_1} = U_N(N_1, O_1) - \lambda_1 + \lambda_2 f'(N_1) = 0 \quad (11)$$

$$\frac{dL}{dN_2} = \gamma U_N(N_2, O_2) - \lambda_2 = 0 \quad (12)$$

$$\frac{dL}{dC_1} = \lambda_1 n'(C_1) - \lambda_4 \beta - \lambda_5 = 0 \quad (13)$$

$$\frac{dL}{dC_2} = \lambda_2 n'(C_2) - \lambda_4 = 0 \quad (14)$$

$$\frac{dL}{dm_1} = -\lambda_3 r P_1 + \lambda_4 \beta + \lambda_5 + \nu_1 = 0 \quad (15)$$

$$\frac{dL}{dm_2} = -\lambda_3 P_2 + \lambda_4 + \nu_2 = 0 \quad (16)$$

### Case 1

If  $\frac{rP_1}{\beta} > P_2$ , equations 15 and 16 imply  $\frac{\lambda_5 + \nu_1}{\beta} > \nu_2$ . Given the non-negativity of the multipliers, it follows that  $\lambda_5 + \nu_1 > 0$ . Since we restrict attention to cases where consumption levels are strictly positive in both periods, we necessarily have  $m_1 > 0$ , thus  $\nu_1 = 0$  and  $\lambda_5 > 0$ . Given that  $\lambda_5 (C_1 - m_1) = 0$ , it

follows that  $C_1 = m_1$  and that the household buys grain in the second period (to maintain  $C_2 > 0$ ):  $m_2 > 0$  (and  $\nu_2 = 0$ ).

Equations 11 and 12 then imply:

$$\frac{U_N(N_1, O_1)}{\gamma U_N(N_2, O_2)} = \frac{\lambda_1 - \lambda_2 f'(N_1)}{\lambda_2}$$

The multipliers  $\lambda_1$  and  $\lambda_2$  can be easily be written as functions of  $\lambda_4$ . Note that 15 and 16 imply  $\lambda_5 = \lambda_4 \left( \frac{rP_1}{P_2} - \beta \right)$ . Then 13 implies  $\lambda_1 = \frac{1}{n'(C_1)} \lambda_4 \frac{rP_1}{P_2}$ . And 14 implies  $\lambda_2 = \frac{\lambda_4}{n'(C_2)}$ . Thus  $\frac{\lambda_1}{\lambda_2} = \frac{n'(C_2)rP_1}{n'(C_1)P_2}$ , and:

$$\frac{U_N(N_1, O_1)}{\gamma U_N(N_2, O_2)} = \frac{\frac{1}{n'(C_1)} rP_1 - \frac{f'(N_1)}{n'(C_2)} P_2}{\frac{P_2}{n'(C_2)}}$$

## Case 2

If  $\frac{rP_1}{\beta} < P_2$ , equations 15 and 16 imply  $\frac{\lambda_5 + \nu_1}{\beta} < \nu_2$ . Given the non-negativity of the multipliers, it follows that  $\nu_2 > 0$  and thus (by complementary slackness)  $m_2 = 0$ . Since we restrict attention to cases where consumption levels are strictly positive in both periods, we necessarily have  $m_1 > 0$  (and thus  $\nu_1 = 0$ ). Furthermore  $C_2 > 0$  implies  $C_1 < y + m_1$ , and thus  $\lambda_5 = 0$ .

Equations 13 and 14 imply  $\lambda_2 = \frac{n'(C_1)}{\beta n'(C_2)} \lambda_1$ . Plugging this expression in 11:  $U_N(N_1, O_1) = \lambda_1 \left( 1 - \frac{n'(C_1)}{\beta n'(C_2)} f'(N_1) \right)$ . Plugging the same expression in 12 implies  $\gamma U_N(N_2, O_2) = \frac{n'(C_1)}{\beta n'(C_2)} \lambda_1$ . And we can write the following expression:

$$\frac{U_N(N_1, O_1)}{\gamma U_N(N_2, O_2)} = \frac{\frac{1}{n'(C_1)} - \frac{1}{\beta n'(C_2)} f'(N_1)}{\frac{1}{\beta n'(C_2)}}$$

## Simulation

We follow Dercon and Krishnan (2000a) and assume that the marginal utility of nutrition has the following functional form:  $U_N(N, O) = N^{-\rho}$ . The marginal utility of the composite good is:  $U_O(N, O) = O^{-\rho}$ . Assuming  $n(c) = c$  and  $f(n) = \varepsilon n$ , the nutrition equations become:  $N_1 = k_1 + C_1$  and  $N_2 = k_2 + \varepsilon C_1 + C_2$ , where  $k_1 = \varepsilon N_0 - a(L_1)$  and  $k_2 = k_1 - a(L_2)$

We consider first the case where  $m_2 > 0$ . Equation 6 implies  $N_1 = N_2 \left( \gamma \frac{rP_1 - \varepsilon P_2}{P_2} \right)^{-\frac{1}{\rho}}$ . We can also write (using 9 and 10):  $O_1 = (\gamma r)^{-\frac{1}{\rho}} O_2$ . Equations 10 and 12 then imply  $\frac{U_O(N_2, O_2)}{U_N(N_2, O_2)} = \frac{\lambda_3}{\lambda_2} = \frac{\lambda_3}{\lambda_2}$ . We can express both  $\lambda_3$  and  $\lambda_2$  as a function of  $\lambda_4$ , using 16 and 14 and we obtain the following expression:  $\frac{U_O(N_2, O_2)}{U_N(N_2, O_2)} = \frac{n'(C_2)}{P_2} = \frac{1}{P_2}$  or  $O_2 = N_2 P_2^{\frac{1}{\rho}}$ . As foodgrain is purchased in both periods, we know  $C_1 = m_1$  and  $C_2 = m_2$ . The budget constraint can thus be written as:  $rP_1 C_1 + rO_1 + P_2 C_2 + O_2 = rR$ . Using



the nutrition equations above, and the various expressions just derived, we obtain an expression of  $N_2$  as a function of the model parameters only:

$$\begin{aligned}
rP_1(N_1 - k_1) + rO_1 + P_2(N_2 - \varepsilon N_1 - k_2 + \varepsilon k_1) + O_2 &= rR \\
N_1(rP_1 - \varepsilon P_2) + rO_1 + P_2 N_2 + O_2 &= rR + k_1(rP_1 - \varepsilon P_2) + k_2 P_2 \\
N_2 \left( \left( \gamma \frac{rP_1 - \varepsilon P_2}{P_2} \right)^{-\frac{1}{\rho}} (rP_1 - \varepsilon P_2) + r(\gamma r)^{-\frac{1}{\rho}} P_2^{\frac{1}{\rho}} + P_2 + P_2^{\frac{1}{\rho}} \right) &= rR + k_1(rP_1 - \varepsilon P_2) + k_2 P_2
\end{aligned}$$

We can then express all other decision variables as a function of these same parameters.

In the second case where  $m_2 = 0$ , equation 7 implies  $N_1 = N_2 \left( \gamma \frac{\beta - \varepsilon}{\beta} \right)^{-\frac{1}{\rho}}$ . Now  $\frac{U_O(N_2, O_2)}{U_N(N_2, O_2)} = \frac{\lambda_4 \frac{\beta}{rP_1}}{\frac{\lambda_4}{n'(C_2)}} = \frac{\beta}{rP_1}$ . Thus  $O_2 = \left( \frac{\beta}{rP_1} \right)^{-\frac{1}{\rho}} N_2$ . The total quantity purchased is  $m_1 = C_1 + \frac{1}{\beta} C_2$ . The budget constraint now is:

$$\begin{aligned}
rP_1 \left( N_1 - k_1 + \frac{1}{\beta} (N_2 - \varepsilon N_1 - k_2 + \varepsilon k_1) \right) + rO_1 + O_2 &= rR \\
N_1 rP_1 \left( 1 - \varepsilon \frac{1}{\beta} \right) + N_2 \frac{rP_1}{\beta} + rO_1 + O_2 &= rR + k_1 rP_1 \left( 1 - \frac{\varepsilon}{\beta} \right) + k_2 \frac{rP_1}{\beta} \\
N_2 \left( \left( \gamma \frac{\beta - \varepsilon}{\beta} \right)^{-\frac{1}{\rho}} rP_1 \left( 1 - \varepsilon \frac{1}{\beta} \right) + \frac{rP_1}{\beta} + r(\gamma r)^{-\frac{1}{\rho}} \left( \frac{rP_1}{\beta} \right)^{\frac{1}{\rho}} + \left( \frac{rP_1}{\beta} \right)^{\frac{1}{\rho}} \right) &= rR + k_1 rP_1 \left( 1 - \frac{\varepsilon}{\beta} \right) + k_2 \frac{rP_1}{\beta}
\end{aligned}$$

Figures 6 and 7 present the effect of a change of  $P_2$  (holding other prices constant) on  $N_1$ ,  $N_2$  and  $C_1 + C_2$ . The parameters are set to the following values:  $r = 1.05$ ,  $\rho = 0.5$ ,  $\gamma = 0.95$ ,  $R = 10$ ,  $k_1 = 0.2$ ,  $k_2 = 0.1$ ,  $\varepsilon = 0.3$ ,  $\beta = 0.9$ . With these parameter values, we are in the case of  $m_2 > 0$  and Figures 6 and 7 obtain.

## Appendix 2

Table 7: Balance tests on baseline characteristics: No-road sample

	2010-2011		2011-2012		2012-2013	
	MEAN	SD	MEAN	SD	MEAN	SD
<b>Panel A : Nutritional Outcomes (Individual level)</b>						
BMI level of adults (19-59 years old)	20.50	2.60	20.13	2.64	20.42	2.64
BMI-for-age z-score of 5-19 years children	-1.00	1.05	-0.98	1.07	-1.08	1.09
BMI-for-age z-score of 0-60 months children	-0.30	1.27	-0.38	1.41	-0.56	1.38
=1 if adults BMI <18.5	0.22	0.41	0.24	0.43	0.22	0.42
=1 if 5-19 years children BMI-for-age <-2	0.17	0.37	0.16	0.37	0.20	0.40
=1 if 0-60 months children BMI-for-age <-2	0.07	0.26	0.09	0.29	0.13	0.33
Difference (after-before rainy season) in BMI level of adults	-0.05	1.36	-	-	-0.83	1.31
Difference in BMI-for-age of 5-19 year children	-0.99	1.06	-	-	-0.82	1.08
Difference in BMI-for-age of 0-60 month children	-0.23	1.43	-	-	-0.36	1.52
<b>Panel B : Foodgrain Production and Market Participation (Household level)</b>						
Cereal production (Kg/cap)	257.80	143.62	109.02	99.59	159.42	118.00
=1 if cereal self-sufficient	0.68	0.47	0.12	0.33	0.41	0.50
=1 if any cereal sale	0.03	0.16	0.06	0.24	0.08	0.27
Cereal sales (Kg/cap)	0.33	2.53	0.98	4.16	1.67	6.41
=1 if any cereal purchase	0.57	0.50	0.85	0.36	0.43	0.50
Cereal purchases (Kg/cap)	13.77	26.45	46.81	45.22	22.44	39.32
=1 if any cereal bulk (>100 kg) purchase	0.39	0.49	0.80	0.40	0.40	0.49
Cereal bulk (>100 kg) purchases (Kg/cap)	13.77	26.45	46.73	45.23	21.97	39.44
<b>Panel C : Foodgrain Purchases (Transaction level)</b>						
=1 if sorghum	0.60	0.49	0.62	0.49	0.42	0.49
Nominal price paid for sorghum	12.79	1.36	19.87	3.45	15.73	2.80
=1 if bought in the village	0.63	0.48	0.25	0.43	0.45	0.50
Nominal price paid for sorghum in the village	13.11	1.28	20.73	3.18	17.18	2.55
=1 if bought during rainy season	0.65	0.48	0.38	0.49	0.46	0.50
Nominal price paid for sorghum during rainy season	12.87	1.36	20.00	4.22	16.65	2.42
=1 if bought before stock depletion	0.57	0.50	0.70	0.46	0.52	0.50
=1 if bought to a particular seller because of - Proximity	0.94	0.24	0.33	0.47	0.45	0.50
=1 if bought to a particular seller because of - Availability	0.03	0.18	0.41	0.49	0.40	0.49
=1 if bought to a particular seller because of - Price	0.03	0.16	0.23	0.42	0.16	0.36

(1) All figures concern the sub-sample of 200 households from control villages.

(2) Foodgrain retail purchases were not investigated in depth at baseline. This explains why bulk purchases correspond to total purchases in 2010-2011.

(3) Panel C includes all bulk purchases of foodgrain and the unit of observation is the 100-kg bag of foodgrain. There are 439 bags of foodgrain bought in 2010-2011, 2076 in 2011-2012 and 670 in 2012-2013.

Table 8: Heterogeneous effects on foodgrain access

	AVAILABILITY		AFFORDABILITY		PURCHASE	
	=1 IF GRAIN BAG BOUGHT LOCALLY	ANNUAL DIST. TRAVELLED PER CAPITA IN MINUTES	PRICE PAID BAG OF SORGHUM IN 1000 CFA-F	=1 IF HH MAKES ANY GRAIN PURCHASE	ANNUAL QUANTITY BOUGHT IN 100KG/CAP <sup>(5)</sup>	ANNUAL EXPENDITURES PER CAPITA IN 1000 CFA-F
	(1)	(2)	(3)	(4)	(5)	(6)
<b>NOROAD = 1 if no road through the village</b>						
AFTER	0.082 [ 0.111]	34.365*** [ 6.081]	5.217*** [ 0.642]	0.540*** [ 0.044]	0.457*** [ 0.065]	7.556*** [ 1.392]
AFTER * TREAT	0.005 [ 0.174]	-7.165 [ 5.354]	-0.586 [ 0.625]	0.034 [ 0.062]	-0.054 [ 0.064]	-0.528 [ 1.313]
AFTER * NOROAD	-0.436*** [ 0.169]	14.195* [ 7.838]	2.613*** [ 0.840]	-0.051 [ 0.111]	-0.035 [ 0.087]	0.334 [ 1.827]
AFTER * TREAT * NOROAD	0.651** [ 0.268]	-8.533 [ 12.160]	-2.445** [ 1.167]	-0.019 [ 0.164]	0.037 [ 0.125]	-1.045 [ 2.295]
LEVEL OF ANALYSIS	BG	HH	BG	HH	HH	HH
OBSERVATION	2516	791	1628	791	791	791

(1) The dependent variables are at bag of grain (BG) or household (HH) level.

(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogeneous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.

(3) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).

(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

(5) Annual foodgrain quantity bought includes retail transactions in addition to bulk purchases.

Table 9: Quantity of foodgrain purchased across quarters

	QUANTITY OF FOODGRAIN PURCHASED			
	OLS		NB <sup>(2)</sup>	
	ALL	NO ROAD	ALL	NO ROAD
	(1)	(2)	(3)	(4)
Q2	0.086*** [ 0.018]	0.091*** [ 0.029]	0.723*** [ 0.179]	0.711** [ 0.335]
Q3	0.111*** [ 0.025]	0.133*** [ 0.041]	0.863*** [ 0.219]	0.923** [ 0.381]
Q4	0.049** [ 0.021]	0.057** [ 0.029]	0.471** [ 0.228]	0.496 [ 0.330]
TREAT	0.021 [ 0.019]	0.030 [ 0.035]	0.167 [ 0.255]	0.150 [ 0.462]
Q2 * TREAT	-0.040 [ 0.026]	-0.058 [ 0.036]	-0.310 [ 0.258]	-0.421 [ 0.419]
Q3 * TREAT	-0.002 [ 0.035]	-0.031 [ 0.068]	-0.069 [ 0.308]	-0.217 [ 0.579]
Q4 * TREAT	-0.032 [ 0.030]	-0.046 [ 0.050]	-0.300 [ 0.318]	-0.397 [ 0.521]
CONTROL GROUP MEAN <sup>(6)</sup>	0.116	0.124	0.116	0.124
OBSERVATION	1580	624	1580	624

(1) Regressions correspond to simple difference estimations and not to differences-in-differences ones.

(2) NB corresponds to negative binomial estimation.

(3) All regressions include control variables corresponding to village-level time-varying exogeneous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.

(4) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).

(5) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

(6) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

Table 10: Heterogeneous effects on foodgrain consumption

	DISPOSABLE		RATION		DIVERSITY	
	LN OF REAL GRAIN DISPOSABLE IN KG/YEAR/CAP	=1 IF REAL GRAIN DISPOSABLE > STANDARD	SELF-REPORTED GRAIN DAILY RATION IN KG/DAY/CAP	=1 IF DAILY RATION CONSIDERED AS SUFFICIENT	HODDINOTT DIETARY DIVERSITY SCORE	IFPRI DIETARY DIVERSITY SCORE
	(1)	(2)	(3)	(4)	(5)	(6)
<b>NOROAD = 1 if no road through the village</b>						
AFTER	-0.410*** [ 0.102 ]	-0.383*** [ 0.082 ]	0.047*** [ 0.006 ]	-0.098* [ 0.057 ]	-	-
AFTER * TREAT	-0.226** [ 0.111 ]	-0.155* [ 0.082 ]	-0.016 [ 0.009 ]	0.045 [ 0.041 ]	-	-
AFTER * NOROAD	-0.188* [ 0.107 ]	-0.111 [ 0.097 ]	-0.009 [ 0.018 ]	-0.081* [ 0.046 ]	-	-
TREAT * NOROAD	0.000 [ 0.000 ]	0.000 [ 0.000 ]	0.000 [ 0.000 ]	0.000 [ 0.000 ]	-14.599 [ 17.207 ]	0.107 [ 0.154 ]
AFTER * TREAT * NOROAD	0.233 [ 0.148 ]	0.193 [ 0.129 ]	0.011 [ 0.021 ]	0.125* [ 0.070 ]	-	-
LEVEL OF ANALYSIS	HH	HH	HH	HH	HH	HH
OBSERVATION	780	780	786	786	393	393

- (1) Real grain disposable corresponds to (production + purchases + gifts in - losses - sales - gifts out).  
(2) The consumption standard in Burkina Faso corresponds to 190 kg/year/capita or, equivalently, 0.520 kg/day/cap.  
(3) Hoddinott (2001) dietary diversity score takes into account all food items as such and the frequency of their consumption (3-level index) over the last month. IFPRI Dietary Diversity Score relies on 9 food categories for its construction and corresponds to the number of food categories consumed over the last day. Here, it has been computed as the number of food categories consumed over the last month.  
(4) All regressions include village fixed effects (except regressions from columns 5 and 6) and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(5) Estimations in column (5) and (6) are simple differences. Baseline data on food diet do not exist.  
(6) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).  
(7) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table 11: Heterogeneous effects on nutrition

	LEVEL				PREVALENCE			
	BMI	BMI-F-AGE Z-SCORE		=1 IF BMI <18.5	=1 IF BMI-F-AGE Z <-2			
	19-59	5-18	3-4	0-2 <sup>(1)</sup>	19-59	5-18	3-4	0-2
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<b>NOROAD = 1 if no road through the village</b>								
AFTER	-0.102 [ 0.124 ]	-0.109** [ 0.052 ]	-0.007 [ 0.126 ]	-0.403** [ 0.169 ]	0.021 [ 0.025 ]	0.043* [ 0.024 ]	0.029 [ 0.028 ]	0.088** [ 0.037 ]
AFTER*TREAT	-0.056 [ 0.175 ]	0.126** [ 0.062 ]	-0.079 [ 0.165 ]	0.331 [ 0.217 ]	0.023 [ 0.027 ]	-0.035 [ 0.035 ]	-0.089** [ 0.036 ]	-0.134** [ 0.059 ]
AFTER*NOROAD	-0.280 [ 0.180 ]	0.033 [ 0.071 ]	0.150 [ 0.185 ]	0.035 [ 0.211 ]	0.008 [ 0.028 ]	-0.037 [ 0.038 ]	-0.128*** [ 0.037 ]	-0.059 [ 0.056 ]
AFTER*TREAT*NOROAD	0.991*** [ 0.253 ]	0.189* [ 0.104 ]	0.095 [ 0.246 ]	0.117 [ 0.327 ]	-0.122*** [ 0.044 ]	-0.014 [ 0.053 ]	0.176*** [ 0.054 ]	0.079 [ 0.078 ]
LEVEL OF ANALYSIS	IND	IND	IND	IND	IND	IND	IND	IND
OBSERVATION	2329	3069	623	721	2329	3069	623	721

- (1) The 0-2 year age category corresponds to children from 6 to 36 months.  
(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(3) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).  
(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table 12: Heterogeneous effects on nutrition for households who did not use FSGs

	LEVEL				PREVALENCE			
	BMI	BMI-F-AGE Z-SCORE			=1 IF BMI <18.5	=1 IF BMI-F-AGE Z <-2		
	19-59	5-18	3-4	0-2 <sup>(1)</sup>	19-59	5-18	3-4	0-2
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
DIFF. AT BASELINE <sup>(5)</sup>	-0.173 [ 0.245 ]	-0.088 [ 0.079 ]	0.053 [ 0.153 ]	-0.184 [ 0.134 ]	0.025 [ 0.033 ]	0.008 [ 0.022 ]	0.015 [ 0.019 ]	0.035 [ 0.023 ]
AFTER (=1 IF POST-TREAT)	-0.064 [ 0.106 ]	-0.088** [ 0.043 ]	0.042 [ 0.125 ]	-0.359** [ 0.142 ]	0.000 [ 0.023 ]	0.030 [ 0.020 ]	-0.001 [ 0.034 ]	0.060* [ 0.032 ]
AFTER + TREAT (DID)	0.249* [ 0.143 ]	0.193*** [ 0.053 ]	-0.093 [ 0.140 ]	0.498*** [ 0.180 ]	-0.025 [ 0.025 ]	-0.047* [ 0.026 ]	-0.024 [ 0.033 ]	-0.084 [ 0.053 ]
CONTROL GROUP MEAN <sup>(6)</sup>	20.575	-1.012	-0.261	-0.513	0.183	0.148	0.044	0.115
LEVEL OF ANALYSIS	IND	IND	IND	IND	IND	IND	IND	IND
OBSERVATION	2072	2696	556	642	2072	2696	556	642

- (1) The 0-2 year age category corresponds to children from 6 to 36 months.  
(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(3) Standard-Errors in brackets are village-level Cluster-Robust-Standard-Errors (CRSE).  
(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .  
(5) Baseline difference in outcomes are computed through separate OLS regressions.  
(6) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

Table 13: Impacts of FSGs on production, income generating activities and visits to the household

	2012 CEREAL PRODUCTION KG/CAP	=1 IF ANY MEMBER INVOLVED IN IGA	TOTAL INCOME FROM IGA 1000 CFA/CAP	ANNUAL CEREAL TRANSFER KG/CAP	=1 IF ANY NON-MEMBER STAY IN IN 2011-2012	NUMBER OF NON-MEMBER STAYING IN MAN/DAY
	(1)	(2)	(3)	(4)	(5)	(6)
TREAT	10.460 [ 14.323 ]	-0.013 [ 0.036 ]	-22.406 [ 35.176 ]	-1.573* [ 0.848 ]	-0.043 [ 0.061 ]	-0.926 [ 0.789 ]
CONTROL GROUP MEAN <sup>(5)</sup>	158.266	0.943	213.887	1.961	0.648	3.510
LEVEL OF ANALYSIS	HH	HH	HH	HH	HH	HH
OBSERVATION	390	391	391	387	386	386

- (1) Estimates correspond to simple difference across treatment status and not to differences-in-differences using baseline.  
(2) All regressions include village fixed effects and control variables corresponding to village-level time-varying exogenous factors : WRSI end-of-season indicator (based on rainfall characteristics) and a set of dummy variables on self-reported agricultural shocks (drought, flood, pest attack). All controls were determined prior to treatment.  
(3) Standard-Errors in brackets are household-level Cluster-Robust-Standard-Errors (CRSE).  
(4) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .  
(5) Control group mean corresponds to the mean of the dependent variable for control villages after treatment.

Table 14: Effects of anticipated purchases on adult nutrition, before and after the lean season

ADULTS BMI LEVEL						
	BEFORE (B)	AFTER (A)	(A)–(B)	BEFORE (B)	AFTER (A)	(A)–(B)
	(1)	(2)	(3)	(4)	(5)	(6)
EFFECT OF =1 IF ANY PURCHASE BEFORE STOCK DEPLETION	0.424*** [ 0.135]	0.063 [ 0.132]	-0.482*** [ 0.172]	0.307** [ 0.139]	-0.060 [ 0.105]	-0.426** [ 0.178]
EFFECT OF QUANTITIES BOUGHT BEFORE STOCK DEPLETION	0.100*** [ 0.029]	0.022 [ 0.023]	-0.094*** [ 0.036]	0.101*** [ 0.034]	-0.005 [ 0.020]	-0.093** [ 0.038]
HH FIXED EFFECTS	YES	YES	YES	NO	NO	NO
IND. FIXED EFFECTS	NO	NO	NO	YES	YES	YES
LEVEL OF ANALYSIS	IND	IND	IND	IND	IND	IND
OBSERVATION	2582	2515	2317	2582	2515	2317

(1) Standard-Errors in brackets are household-level Cluster-Robust-Standard-Errors (CRSE).  
(2) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table 15: Daily foodgrain ration and the timing of purchase using monthly data

CEREAL DAILY RATION IN KG/CAP				
	P = 1 IF ANY CEREAL PURCHASE		P = CEREAL QUANTITIES PURCHASED IN KG/CAP	
	(1)	(2)	(3)	(4)
FEB	-0.052 [ 0.040]	-0.039 [ 0.036]	-0.018 [ 0.035]	-0.021 [ 0.032]
MAR	-0.121*** [ 0.045]	-0.091*** [ 0.034]	-0.097** [ 0.043]	-0.072** [ 0.032]
APR	-0.105** [ 0.045]	-0.087** [ 0.040]	-0.078** [ 0.037]	-0.074** [ 0.036]
MAY	-0.083** [ 0.039]	-0.075** [ 0.037]	-0.041 [ 0.034]	-0.051 [ 0.035]
JUN	-0.098** [ 0.038]	-0.080** [ 0.036]	-0.068** [ 0.033]	-0.060* [ 0.034]
JUL	-0.000 [ 0.032]	0.016 [ 0.030]	0.039 [ 0.029]	0.036 [ 0.028]
AUG	0.011 [ 0.041]	0.029 [ 0.038]	0.034 [ 0.034]	0.041 [ 0.035]
SEP	0.028 [ 0.036]	0.045 [ 0.034]	0.042 [ 0.032]	0.052* [ 0.031]
OCT	-0.076* [ 0.040]	-0.055 [ 0.038]	-0.042 [ 0.034]	-0.030 [ 0.034]
P	-0.143*** [ 0.053]	-0.078* [ 0.042]	-0.002 [ 0.001]	-0.001 [ 0.001]
FEB * P	0.121* [ 0.068]	0.110* [ 0.058]	-0.000 [ 0.004]	0.005 [ 0.004]
MAR * P	0.209 [ 0.135]	0.102 [ 0.065]	0.010 [ 0.007]	0.002 [ 0.002]
APR * P	0.125* [ 0.068]	0.090* [ 0.053]	0.004 [ 0.003]	0.004 [ 0.003]
MAY * P	0.197** [ 0.097]	0.194** [ 0.097]	0.004 [ 0.007]	0.011 [ 0.008]
JUN * P	0.165** [ 0.066]	0.102 [ 0.065]	0.005 [ 0.004]	0.002 [ 0.005]
JUL * P	0.115** [ 0.055]	0.057 [ 0.047]	-0.002 [ 0.003]	-0.002 [ 0.002]
AUG * P	0.126* [ 0.068]	0.059 [ 0.062]	0.005* [ 0.003]	0.002 [ 0.003]
SEP * P	0.123* [ 0.064]	0.052 [ 0.056]	0.006 [ 0.004]	0.002 [ 0.003]
OCT * P	0.188** [ 0.073]	0.112* [ 0.062]	0.009 [ 0.006]	0.001 [ 0.006]
CEREAL DISPOSABLE IN KG/CAP	0.001*** [ 0.000]	-	0.001*** [ 0.000]	-
FE HH CRSE HH	NO YES	YES YES	NO YES	YES YES
LEVEL OF ANALYSIS OBSERVATION	MONTH 726	MONTH 726	MONTH 726	MONTH 726

(1) Level of significance : \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ 

(2) HH-level Cluster-Robust-Standard-Errors in brackets.



### Appendix 3

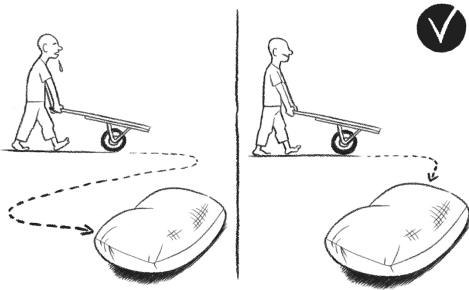


Figure 12: Visual aid example “Households bought foodgrain at lower prices in FSG villages”

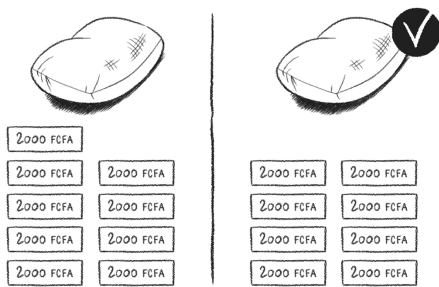


Figure 13: Visual aid example “Households did not consume more or better food in FSG villages”



Figure 14: Visual aid example “Households had a better nutritional status after the drought in FSG villages”

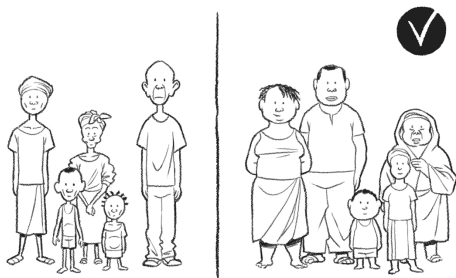


Figure 15: Visual aid example “The paradox”

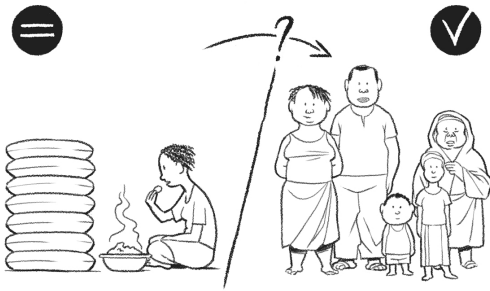


Figure 16: Canonical patterns of timing of purchases

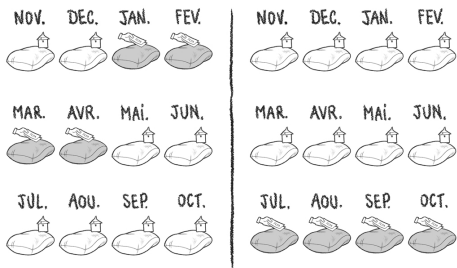


Figure 17: Canonical patterns of timing of purchases and consumption

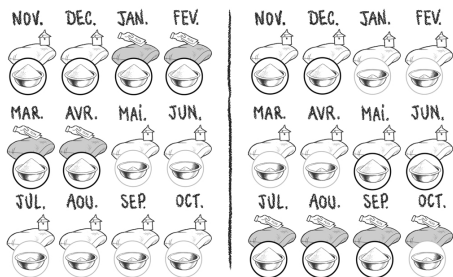


Figure 18: Example of board and cards

